Statistiek II

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We often wish to compare two different variables

Examples: compare results on two distinct tests

- age and ability
- education (in years) and income
- speed and accuracy

Methods to compare two (or more) variables:

- Correlation coefficient
- Regression analysis

Notice:

Correlation and regression only for numeric variables!

Terminology: we speak of

- cases, e.g., Joe, Sam, etc. and
- ▶ variables, e.g., height (*h*) and weight (*w*)
- ► Then each variable has a value for each case; h_j is Joe's height, and w_s is Sam's weight

We compare two variables by comparing their values for a set of cases:

- h_j versus w_j
- ▶ h_s versus w_s
- etc.

Example: Hoppenbrouwers measured pronunciation differences among pairs of dialects. We compare these to the geographic distance between places where they are spoken.

Dialect pair	Phon. dist.	Geogr. dist.
Almelo/Haarlem	0.58	100
Almelo/Kerkrade	1.18	200
Almelo/Makkum	0.90	250
Almelo/Roodeschool	0.81	220
Almelo/Soest	0.91	70
Haarlem/Kerkrad	1.06	230
:	:	:
Kerkrade/Soest	1.14	201
Makkum/Rodeschool	0.95	125
Makkum/Soest	1.00	216
Roodeschool/Soest	0.94	163

Two variables—phonetic and geographic distance, and 15 cases (here, each pair is a separate case)

One useful technique is to visualize the relation by graphing it:



Scatterplot shows the relationship between two quantitative variables

Scatterplots

Each dot is a case, whose *x*-value is geographic distance, and *y*-value is phonetic distance.



In general, we use x-axis for **independent** variables, and y-axis for **dependent** ones. We don't know whether phonetic distance depends on geographic distance, but it might (while reverse is implausible).

Least squares regression

The simplest form of dependence is **linear**—the independent variable determines a portion of the dependent value.

We can visualize this by fitting a straight line to the scatterplot:



If the scatterplot clearly suggests not a straight line, but rather a curve of another sort, you probably need to first **transform** one of the data sets.

This is an advanced topic, but something to keep in mind!

Least squares regression



Like every straight line, this has an equation of the form: y = a + bx

a is the point where the line crosses the *y*-axis, the **intercept**, and b the **slope**.

Predicted vs observed values

The independent variable determines the dependent value (somewhat); this is the predicted value \hat{y} —the value on the line.

Note also that the actual value y—the data dot—is not always the same as \hat{y}



The difference between observed and predicted values

$$\epsilon_i := (y_i - \hat{y}_i)$$

is the **residual**—what the linear model does not predict. It is the vertical distance between the data point and the regression line.

Least-squares regression finds the line which minimizes the squared residuals—for all the data:

$$\sum_{i=1}^{n} \epsilon_i^2 = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

Least squares regression finds the best straight line which models the data (minimizes the squared error).

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Call:

lm(formula = phonetic distance \sim geographic distance)

Coefficients:

(Intercept) geographic distance

0.653292 0.001618
```

Regression line: y = 0.65 + 0.0016x

Residuals

Regression finds the best line, but is sensitive to extreme values. Examine residuals.



Note: requirement in regression model that residuals be normally distributed. Check with normal QQ-plot!

Check normality of residuals



Normal Q-Q Plot: residuals

Residuals look reasonably normal (Shapiro-Wilk test p = 0.18)

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R plot of residuals



Save residuals as new variable, then graph against original x value

Watch out for extreme x values—influential, though residual may be small. See example 2.12 in Moore & McCabe.

Also examine **outliers**—large residuals.

How does regression work?

Suppose we have a sample $S = (x_i, y_i)$ with i = 1, ..., n.

Let
$$x := (x_1, ..., x_n)$$
 and $y := (y_1, ..., y_n)$

We want to **estimate the regression line** y = a + bx for this data.

This amounts to optimizing the intercept a and slope b with respect to the residuals:

Find a and b such that for a given sample S the sum of squared residuals is minimized.

We express the sum of squared residuals as a function of the (unknown) regression line:

$$\sum_{i=1}^{n} \epsilon_{i}^{2} = \sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}$$

$$= \sum_{i=1}^{n} (y_{i} - (a + bx_{i}))^{2}$$

$$= \sum_{i=1}^{n} (y_{i} - a - bx_{i})^{2}$$

$$= \sum_{i=1}^{n} (a^{2} + 2abx_{i} - 2ay_{i} + b^{2}x_{i}^{2} - 2bx_{i}y_{i} + y_{i}^{2})$$

Thus, $\sum_{i=1}^{n} \epsilon_i^2$ is function f in x, y with unknown parameters a, b.

Estimating the regression line

For a fixed sample S = (x, y), we want to minimize $f_{ab}(x, y)$ with

$$f_{ab}(x,y) = \sum_{i=1}^{n} (a^2 + 2abx_i - 2ay_i + b^2x_i^2 - 2bx_iy_i + y_i^2)$$

To minimize this function, find a and b such that $f'_{ab}(x, y) = 0$.

Treat a and b as variables and find partial derivatives $\frac{\partial}{\partial a}f$, $\frac{\partial}{\partial b}f$

$$\frac{\partial}{\partial a}f = f'_{xyb}(a) = \sum_{i=1}^{n} (2a + 2bx_i - 2y_i)$$
$$\frac{\partial}{\partial b}f = f'_{xya}(b) = \sum_{i=1}^{n} (2ax_i + 2bx_i^2 - 2x_iy_i)$$

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Dialect pair	Phon. dist.	Geogr. dist.
Almelo/Haarlem	0.58	100
Almelo/Kerkrade	1.18	200
Kerkrade/Roodeschool	1.27	300

- plug these sample values into partial derivatives
- set them to zero
- solve pair of linear equations

$$f'_{xyb}(a) = \sum_{i=1}^{n} (2a + 2bx_i - 2y_i)$$

= $2a + 2b \cdot 100 - 2 \cdot 0.58 + 2a + 2b \cdot 200 - 2 \cdot 1.18 + 2a + 2b \cdot 300 - 2 \cdot 1.27$
= $6a + 1200b - 6.06$

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Regression—tiny example

$$f'_{xya}(b) = \sum_{i=1}^{n} (2ax_i + 2bx_i^2 - 2x_iy_i)$$

= $2a \cdot 100 - 2b \cdot (100)^2 - 2 \cdot 100 \cdot 0.58 + 2a \cdot 200 - 2b \cdot (200)^2 - 2 \cdot 200 \cdot 1.18 + 2a \cdot 300 - 2b \cdot (300)^2 - 2 \cdot 300 \cdot 1.27$
= $1200a + 280.000b - 1350$

Set to zero and solve:

$$0 = 6a + 1200b - 6.06$$
(I)

$$\Leftrightarrow \quad 0 = a + 200b - 1.01$$

$$\Leftrightarrow \quad a = 1.01 - 200b$$

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Regression—tiny example

$$a = 1.01 - 200b$$
 (I)
 $0 = 1200a + 280.000b - 1350$ (II)

Substitute a in (II) by (I):

 $0 = 1200 \cdot (1.01 - 200b) + 280.000b - 1350$ $\Leftrightarrow \quad 0 = 1212 - 240.000b + 280.000b - 1350$ $\Leftrightarrow \quad 40.000b = 1350 - 1212$ $\Leftrightarrow \quad b = \frac{138}{40.000} = \underline{0.00345}$ $\Rightarrow \quad a = 1.01 - 200 \cdot 0.00345 = \underline{0.32}$

Hence, the regression line y = 0.32 + 0.00345x minimizes the sum of squared residuals

Check calculations with R



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- Regression is asymmetric—appropriate when one variable might be 'explained' by a second
 - Reading times on the basis of difficulty—negative!
 - Child's ability on the basis of parents' ability
 - Final grade based on class attendance, etc.
- No answer (yet) to how well does x explain y Correlation analysis provides an answer
- Correlation symmetric measure of extent to which variables predict each other
- Answer to how well does x explain y

Regression and correlation inappropriate when 'best fit' is not straight line (need data transformations)

How do you know if you are going to do well in a stats course?

Suppose you spend a lot of time on the material—more than your average class mate—then you'll have a high z-score in the distribution of study time.

You know that, generally, study time predicts grades.

So you know that you should have a high z-score in the distribution of grades.

If your final grade is not so good, I would expect you didn't spend much time studying. You would be below the mean in both distributions and have negative z-scores.

Correlation coefficient

If $x = (x_1, \ldots, x_n)$ is study time, and $y = (y_1, \ldots, y_n)$ are grades, we can measure correlation between the two variables as

$$r_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} z_{x_i} \cdot z_{y_i}$$

- compute everyone's z-score (study time and grades)
- multiply both z-scores and sum for everyone in class
- divide by the degrees of freedom (# students -1)

Note: positive sum results from multiplying two positive or negative z-scores for x and y (positive correlation)

Negative sum (correlation) results from multiplying positive and negative z-scores (and vice versa)

No correlation results from mixed-sign z-scores with sum close to zero.

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Correlation coefficient aka "Pearson's product-moment coefficient"

$$r_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} z_{x_i} \cdot z_{y_i} = \frac{1}{n-1} \sum_{i=1}^{n} \left(\frac{x_i - \overline{x}}{\sigma_x} \right) \left(\frac{y_i - \overline{y}}{\sigma_y} \right)$$

• r_{xy} reflects the strength of the relation between x and y

•
$$r_{xy} = 0$$
 no correlation

 r_{xy} = 1 perfect positive correlation (all data points on a straight line with positive slope)

•
$$r_{xy} = -1$$
 perfect negative correlation

- no necessary dependence!
 - shoe size and reading ability correlate—both dependent on age

Visualizing correlation



- data points lie close to the regression line
- correlation coefficient $r_{xy} = 0.83$
- strong positive correlation

Visualizing correlation



- data points scatter in a cloud around regression line
- correlation coefficient $r_{xy} = 0.1$
- no correlation (there might be correlation in both subsets)

Visualizing correlation



- data points close to regression line with negative slope
- correlation coefficient $r_{xy} = -0.77$
- correlation, but negative

In our example: correlation coefficient for geographic and phonetic distance

In R simply call:

cor(phonetic-distance,geographic-distance)
[1] 0.6574452

Hence, phonetic and geographic distance correlate at r = 0.66

r is a 'plain number'—no units

 insensitive to scale, percentages, etc.
 E.g., correlation with temperature can ignore scale (Celsius vs Fahrenheit)

• symmetric $r_{xy} = r_{yx}$

Properties of correlation

$$r_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} \left(\frac{x_i - \overline{x}}{\sigma_x} \right) \left(\frac{y_i - \overline{y}}{\sigma_y} \right)$$

- correlation requires that both variables be quantitative (numerical)
- \blacktriangleright correlation coefficient always between 1 and -1
- \blacktriangleright as $r \rightarrow 1$ (or -1), dots cluster near regression line
- r measures 'clustering' relative to standard deviations σ_x , σ_y
- correlation can be misleading in the presence of outliers or nonlinear association
- therefore...

...always plot your data



Four variables y have same mean, standard deviation, correlation and regression line (examples from Anscombe)

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Relationship between correlation and regression

Recall we obtained two partial derivatives (when minimizing sum of squared residuals):

$$f'_{xyb}(a) = \sum_{i=1}^{n} (2a + 2bx_i - 2y_i)$$
(1)
$$f'_{xya}(b) = \sum_{i=1}^{n} (2ax_i + 2bx_i^2 - 2x_iy_i)$$
(2)

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Set (1) to zero:

$$f'_{xyb}(a) = 0$$

$$\Leftrightarrow \quad n \cdot 2a + \sum_{i=1}^{n} (2bx_i - 2y_i) = 0$$

$$\Leftrightarrow \quad n \cdot 2a + 2b \sum_{i=1}^{n} x_i - 2 \sum_{i=1}^{n} y_i = 0$$

$$\Leftrightarrow \quad n \cdot a = n \cdot \overline{y} - n \cdot b\overline{x}$$

$$\Leftrightarrow \quad a = \overline{y} - b\overline{x}$$

Relationship between correlation and regression

Plug $a = \overline{y} - b\overline{x}$ into (2) and set to zero:

$$f'_{xya}(b) = 0$$

$$\Leftrightarrow \sum_{i=1}^{n} (2(\overline{y} - b\overline{x})x_i + 2bx_i^2 - 2x_iy_i) = 0$$

$$\Leftrightarrow (\overline{y} - b\overline{x})(n\overline{x}) + b\sum_{i=1}^{n} x_i^2 - \sum_{i=1}^{n} x_iy_i = 0$$

$$\Leftrightarrow n\overline{x}\overline{y} - b\overline{x}^2n + b\sum_{i=1}^{n} x_i^2 - \sum_{i=1}^{n} x_iy_i = 0$$

$$\Leftrightarrow b(\sum_{i=1}^{n} x_i^2 - \overline{x}^2n) = \sum_{i=1}^{n} x_iy_i - n\overline{x}\overline{y}$$

$$\Leftrightarrow b = \frac{\sum_{i=1}^{n} x_iy_i - n\overline{x}\overline{y}}{\sum_{i=1}^{n} x_i^2 - \overline{x}^2n}$$

Relationship between correlation and regression

$$b = \frac{\sum_{i=1}^{n} x_i y_i - n\overline{x}\overline{y}}{\sum_{i=1}^{n} x_i^2 - \overline{x}^2 n} \quad \Leftrightarrow \quad b = \frac{\sum_{i=1}^{n} x_i y_i - n\overline{x}\overline{y}}{\sum_{i=1}^{n} (x_i - \overline{x})^2}$$
$$\Leftrightarrow \quad b = \frac{\sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y})}{\sum_{i=1}^{n} (x_i - \overline{x})^2}$$
$$\Leftrightarrow \quad b = \frac{1}{n-1} \frac{\sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y})}{\left(\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \overline{x})^2\right)}$$
$$\Leftrightarrow \quad b = \frac{1}{n-1} \sum_{i=1}^{n} \frac{(x_i - \overline{x})(y_i - \overline{y})}{\sigma_x^2}$$
$$\Leftrightarrow \quad b = \left(\frac{1}{n-1} \sum_{i=1}^{n} \left(\frac{x_i - \overline{x}}{\sigma_x}\right) \left(\frac{y_i - \overline{y}}{\sigma_y}\right)\right) \cdot \frac{\sigma_y}{\sigma_x}$$
$$\Leftrightarrow \quad b = r \frac{\sigma_y}{\sigma_x}$$

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Thus, the regression line y = a + bx has

- ▶ slope $b = r \frac{\sigma_y}{\sigma_y}$ and
- intercept $a = \overline{y} b\overline{x}$

Consequently:

- correlation and regression are related via the coefficient r
- ▶ regression line always flatter than SD line, the line with slope $\frac{\sigma_y}{\sigma_x}$ which passes through $(\overline{x}, \overline{y})$

What's the point of regression analysis?

- analyze y as dependent on x (non-symmetric)
- determine how much of y's variance can be attributed to x

Correlation and regression



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As in ANOVA, we can partition the variance in regression model:

$$\sum_{i=1}^{n} (y_i - \overline{y})^2 = \sum_{i=1}^{n} (y_i - \underbrace{(a + bx_i)}_{\text{regression line}})^2 + \sum_{i=1}^{n} (\underbrace{(a + bx_i)}_{\text{regression line}} - \overline{y})^2$$

Total variance = Unexplained variance + Explained variance

To what extent does explanatory variable x explain variation in response y? The quotient

$$\frac{\sum_{i=1}^{n} ((a + bx_i) - \overline{y})^2}{\sum_{i=1}^{n} (y_i - \overline{y})^2} = \frac{\text{explained variance}}{\text{total variance}}$$

measures this precisely.

Another relation between correlation and regression

 $\frac{\text{explained variance}}{\text{total variance}}$

$$= \frac{\sum_{i=1}^{n} ((a + bx_i) - \overline{y})^2}{\sum_{i=1}^{n} (y_i - \overline{y})^2}$$

$$= \frac{\sum_{i=1}^{n} ((\overline{y} - b\overline{x} + bx_i) - \overline{y})^2}{\sum_{i=1}^{n} (y_i - \overline{y})^2}$$

$$= \frac{\sum_{i=1}^{n} b^2 (x_i - \overline{x})^2}{\sum_{i=1}^{n} (y_i - \overline{y})^2}$$

$$= b^2 \cdot \left(\frac{\sigma_x}{\sigma_y}\right)^2$$

$$= r^2 \left(\frac{\sigma_y}{\sigma_x}\right)^2 \cdot \left(\frac{\sigma_x}{\sigma_y}\right)^2$$

$$= r^2$$

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 $\frac{\text{explained variance}}{\text{total variance}} = r^2 \qquad (\text{``coefficient of determination''})$

- r² indicates proportion of variability in data set that is accounted for by regression model
- provides a measure of how well future outcomes are likely to be predicted by the model
- in our example (phonetic distance of dialects):

$$r^2 = 0.66^2 = 0.435$$

Thus, 44% of the phonetic variation between dialects is accounted for by geographic distance

Interpretation of correlation via averages

Example: height, weight have correlation coefficient $r_{hw} = 0.5$

$$\mu_h = 178$$
cm, $\mu_w = 72$ kg, $\sigma_h = 6$ cm, $\sigma_w = 6$ kg

Slope of regression line: $b = r \cdot \frac{\sigma_w}{\sigma_h}$, i.e., for every σ_h , predicted weight changes by $r \cdot \sigma_w$

What is the average weight of those 184cm tall?

$$184\text{cm} = 178\text{cm} + 6\text{cm}$$

$$= \mu_h + 1 \cdot \sigma_h$$

$$\delta_{\sigma_h} = 1$$

$$\overline{w}_{184\text{cm}} = \mu_w + r_{hw} \cdot \delta_{\sigma_h} \cdot \sigma_w$$

$$= 72\text{kg} + 0.5 \cdot 1 \cdot 6\text{kg} = 75\text{kg}$$

In regression, for each σ_x , the predicted value of y changes by $r\sigma_y$ When there is less than perfect correlation, $0 \le r < 1$ Hence, a predicted z_y for y will be closer to (the mean) 0 than z_x

In the previous example:

$$z_{\rm x} = rac{184_{
m cm} - 178_{
m cm}}{6_{
m cm}} = 1$$
, $z_{y} = rac{75_{
m kg} - 72_{
m kg}}{6_{
m kg}} = 0.5$

Since r < 1, averages of correlated variables **must** regress toward the mean $(z_y = r \cdot z_x)$

Regression toward the mean is a mathematical inevitability

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Regression fallacy: seeing causation in regression

Examples:

(1) height correlation between parents and children (r = 0.4) due to regression toward the mean, very tall parents tend to have less tall children (still taller than average)

Regression fallacy: tall father concludes his wife must have cheated

(2) motivation correlates with exam scores (r = 0.5)

test-retest situations show extremes (high and low scores) closer to mean on second test (regression toward mean)

Regression fallacy: bad students improved because I punished them

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Correlation

Properties:

- only for numeric variables
- measures strength of a linear relation
- symmetric $r_{xy} = r_{yx}$
- related to the slope of the regression line

Caution needed:

- non-linear associations, i.e., curved patterns
- individual points with large residuals (outliers)
- influential observations (large deviation in x direction)
- "ecological correlations", i.e., correlations based on averages, popular in politics, overstate size of r
- correlation \neq causation (e.g., shoe size and reading ability)

Inference for regression

Test whether regression yields significant association of variables:

Residual standard error: estimated standard error about the regression line

$$s = \sqrt{\frac{\sum_{i=1}^{n} e_i^2}{n-2}}$$

Standard error of the regression slope:

$$\mathsf{SE}_b = rac{s}{\sqrt{\sum_i^n (x_i - \overline{x})^2}}$$

We test:
$$H_0: b = 0, \quad H_a: b \neq 0$$

Calculate *t*-statistic: $t = \frac{b}{SE_b}$

Compare with critical t^* from t(n-2)

In our example (phonetic variation in dialects):

$$s = \sqrt{\frac{0.3056}{13}} = 0.1533$$

$$\mathsf{SE}_b = \frac{0.1533}{\sqrt{298.2}} = 0.000514$$

$$t = \frac{0.001618}{0.000514} = 3.148$$

Critical value $t^* = 2.16$ (for t(df=13), $\alpha = 0.05$), hence reject H_0 :

The data provides evidence in favor of a relationship between geographic and phonetic distance

Call: $Im(formula = phonetic distance \sim geographic distance)$							
Residuals:	N.4.	10	N 4 1'	20			
	Min -0.2496	1Q -0.1015	0.0288	3Q 0.1129	Max 0.2032		
Coefficients:							
	Estimate	Std. Error	t value	$\Pr(> t)$			
(Intercept)	0.653292	0.104245	6.27	2.9e-05 ***			
geographic distance	0.001618	0.000514	3.15	0.0077 **			
Signif. codes:	0 ***	0.001 **	0.01 *	0.05 .			
Residual standard error: Multiple R-Squared: F-statistic:	0.153 on 13 degrees of freedom 0.432, Adjusted R-squared: 0.389 9.9 on 1 and 13 DF, p-value: 0.00773						

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Confidence intervals

What is the mean phonetic distance of dialects for $x^* = 150$ km geographic distance?

$$\hat{y} = 0.65 + 0.0016 \cdot 150 = 0.89$$

Standard error for mean response \hat{y} (for fixed x^*):

$$\begin{aligned} \mathsf{SE}_{\hat{y}} &= s \cdot \sqrt{\frac{1}{n} + \frac{(x^* - \overline{x})^2}{\sum_i^n (x_i - \overline{x})^2}} \end{aligned}$$

Here:
$$\begin{aligned} \mathsf{SE}_{\hat{y}} &= 0.1533 \cdot \sqrt{\frac{1}{15} + \frac{(150 - 187.5)^2}{88914}} = 0.04403 \end{aligned}$$

Confidence: $\hat{y} \pm t^* SE_{\hat{y}} = 0.89 \pm 2.16 * 0.04403 = 0.89 \pm 0.0951$

Hence, with 95% certainty, mean phonetic distance (for $x^* = 150$ km) lies in the interval Cl=(0.795,0.985)

Visualizing confidence intervals

Next week: multiple regression

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