

Statistiek II

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- 1 One-way ANOVA.
- 2 Factorial ANOVA.
- 3 Repeated measures ANOVA.
- 4 Correlation and regression.
- 5 Multiple regression.
- 6 Logistic regression.
- 7 Hierarchical, or “mixed” models

Today: One-way ANOVA

- 1 General motivation
- 2 F -test and F -distribution
- 3 ANOVA example
- 4 The logic of ANOVA

Short break

- 5 ANOVA calculations
- 6 Post-hoc tests

What's ANalysis Of VAriance (ANOVA)?

- ▶ Most popular statistical test for numerical data
- ▶ Generalized t -test
- ▶ Compares means of more than two groups
- ▶ Fairly robust
- ▶ Based on F -distribution
- ▶ compares variances (between groups and within groups)
- ▶ Two basic versions:
 - a One-way (or single) ANOVA: compare groups along one dimension, e.g., grade point average by school class
 - b N-way (or factorial) ANOVA: compare groups along ≥ 2 dimensions, e.g., grade point average by school class and gender

Typical applications

- ▶ One-way ANOVA:
Compare time needed for lexical recognition in
 1. healthy adults
 2. patients with Wernicke's aphasia
 3. patients with Broca's aphasia

- ▶ Factorial ANOVA:
Compare lexical recognition time in male and female in the same three groups

Comparing multiple means

- ▶ For **two** groups: use t -test
- ▶ Note: testing for p-value of 0.05 shows significance 1 time in 20 if there is no difference in population mean (effect of chance)
- ▶ But suppose there are 7 groups, i.e., we test $\binom{7}{2} = 21$ pairs of groups
- ▶ **Caution:** several tests (on the same data) run the risk of finding significance through sheer chance

Multiple comparison problem

Example: Suppose you run $k = 3$ tests, always seeking a result significant at $\alpha = 0.05$

⇒ probability of getting at least one false positive is given by:

$$\begin{aligned}\alpha_{FW} &= 1 - P(\text{zero false positive results}) \\ &= 1 - (1 - \alpha)^k \\ &= 1 - (1 - 0.05)^3 \\ &= 1 - (0.95)^3 \\ &= 0.143\end{aligned}$$

Hence, with only 3 pairwise tests, the chance of committing type I error almost 15% (and 66% for 21 tests!)

α_{FW} called Bonferroni family-wise α -level

Bonferroni correction for multiple comparisons

To guarantee a **family-wise** α -level of 0.05, divide α by number of tests.

Example: $0.05/3$ ($= \alpha/\#$ tests) $= 0.017$ (note: $0.983^3 \approx 0.95$)
 \Rightarrow set $\alpha = 0.017$ ($=$ Bonferroni-corrected α -level)

- ▶ If p-value is less than the Bonferroni-corrected target α : reject the null hypothesis.
- ▶ If p-value greater than the Bonferroni-corrected target α : do not reject the null hypothesis.

Analysis of variance

- ▶ ANOVA automatically corrects for looking at several relationships (like Bonferroni correction)
- ▶ Based on F -distribution: Moore & McCabe, §7.3, pp. 435–445
- ▶ Measures the difference between two variances (variance σ^2)

$$F = \frac{s_1^2}{s_2^2}$$

- ▶ always positive since variances are positive
- ▶ two degrees of freedom interesting, one for s_1 , one for s_2

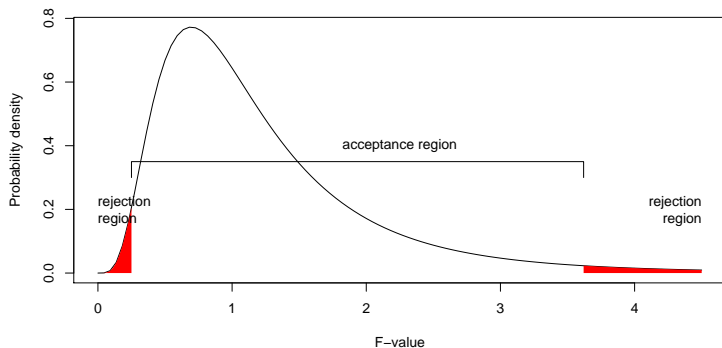
F-test vs. F-distribution

F-value: $F = \frac{s_1^2}{s_2^2}$

- ▶ F-values used in F-test (Fisher's test)
 - H_0 : samples are from same distribution ($s_1 = s_2$)
 - H_a : samples are from different distributions ($s_1 \neq s_2$)
 - value near 1 indicates same variance
 - value near 0 or $+\infty$ indicates difference in variance
- ▶ F-test very sensitive to deviations from normal
- ▶ ANOVA uses F-distribution, but is different: ANOVA \neq F-test!

F-distribution

Critical area for F -distribution at $p = 0.05$ (df: 12,10)



Note the symmetry: $P\left(\frac{s_1^2}{s_2^2} < x\right) = P\left(\frac{s_2^2}{s_1^2} > \frac{1}{x}\right)$

(because $y < x \Leftrightarrow \frac{1}{y} > \frac{1}{x}$ for $x, y \in \mathbb{R}^+$)

Example: height

group	sample size	mean	standard deviation
boys	16	180cm	6cm
girls	9	168cm	4cm

Is the difference in standard deviation significant?

Examine $F = \frac{s_{\text{boys}}^2}{s_{\text{girls}}^2}$

Degrees of freedom: $df_{\text{boys}} = 16 - 1$
 $df_{\text{girls}} = 9 - 1$

F-test critical area (for two-tailed test with $\alpha = 0.05$)

$$P(F(15, 8) > x) = \frac{\alpha}{2} = 0.025$$

$$P(F(15, 8) < x) = 1 - 0.025$$

$$P(F(15, 8) < \underline{4.1}) = 0.975 \quad \text{Moore \& McCabe, Table E, p. 706}$$

(no values directly for $P(F(df_1, df_2) > x)$)

$$P(F(15, 8) < x) = 0.025$$

$$\Leftrightarrow P(F(8, 15) > x') = 0.025 \quad \text{where } x' = \frac{1}{x}$$

$$\Leftrightarrow P(F(8, 15) > 3.2) = 0.025 \quad (\text{tables})$$

$$\Leftrightarrow P(F(15, 8) < \frac{1}{3.2}) = 0.025$$

$$\Leftrightarrow P(F(15, 8) < \underline{0.31}) = 0.025$$

Reject H_0 if $F < 0.31$ or $F > 4.1$

Here, $F = \frac{6^2}{4^2} = 2.25$ (hence no evidence of difference in distributions)

Analysis of Variance (ANOVA) most popular statistical test for numerical data

- ▶ several types
 - single, “one-way”
 - factorial, “two-, three-, . . . , n-way”
 - single/factorial repeated measures
- ▶ examines variation
 - “between-groups”—gender, age, etc.
 - “within-groups”—overall
- ▶ automatically corrects for looking at several relationships (like Bonferroni correction)
- ▶ uses F -distribution, where $F(n, m)$ fixes n typically at the number of groups (minus 1), m at the number of subjects, i.e., data points (minus number of groups)

Detailed example: one-way ANOVA

Question: Are exam grades of **four** groups of foreign students “Nederlands voor anderstaligen” the same? More exactly, are the four averages the same?

$$H_0 : \mu_1 = \mu_2 = \mu_3 = \mu_4$$

$$H_a : \mu_1 \neq \mu_2 \text{ or } \mu_1 \neq \mu_3 \dots \text{ or } \mu_3 \neq \mu_4$$

Alternative hypothesis: at least one group has a different mean

For the question of whether any particular pair is different, the t -test is appropriate.

For testing whether all language groups are the same, pairwise t -tests *exaggerate* differences (increase the chance of type I error)

We therefore want to apply one-way ANOVA

Data: Dutch proficiency of foreigners

Four groups of ten students each:

	Group			
	Europe	America	Africa	Asia
	10	33	26	26
	19	21	25	21
	⋮	⋮	⋮	⋮
	31	20	15	21
Mean	25.0	21.9	23.1	21.3
Samp. SD	8.14	6.61	5.92	6.90
Samp. Variance	66.22	43.66	34.99	47.57

ANOVA conditions

ANOVA assumptions:

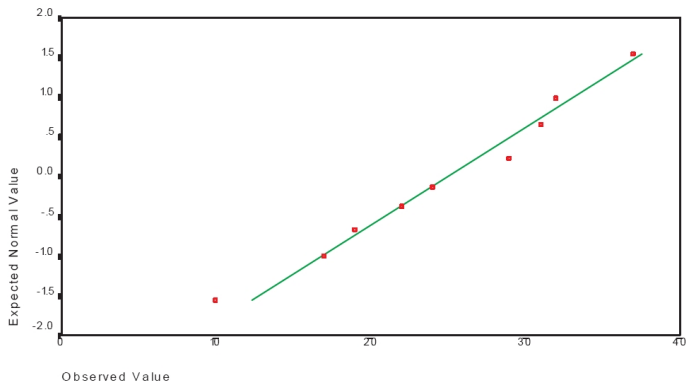
- ▶ Normal distribution per subgroup
- ▶ Same variance in subgroups: least SD $>$ one-half of largest SD
- ▶ **independent** observations: watch out for test-retest situations!

Check differences in SD's! (some SPSS computing)

Variable	Std Dev	Valid	
		N	Label
Europa	8.14	10	
America	6.61	10	
Africa	5.92	10	
Azie	6.90	10	

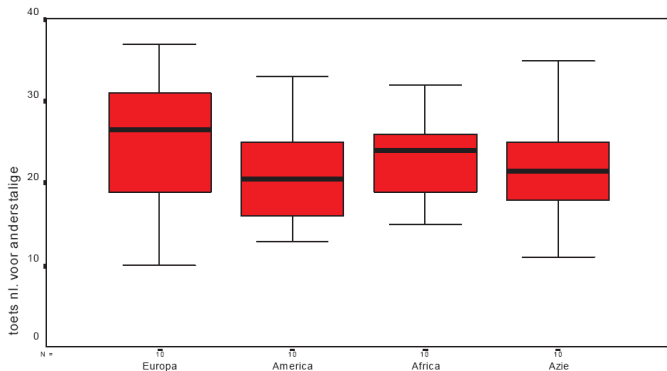
Assumption: normal distribution **per group**, check with normal quantile plot, e.g., for Europeans below (repeat for every group)

Normal Q-Q plot of toets.nl voor anderstalige



Visualizing ANOVA data

Is there a significant difference in the means (of the groups being contrasted)?



Take care that boxplots sketch **medians** not **means**.

Sketch of ANOVA

Group			
1	2	3	4
Eur.	Amer.	Africa	Asia
\vdots	\vdots	\vdots	\vdots
x_{1j}	x_{2j}	x_{3j}	x_{4j}
\vdots	\vdots	\vdots	\vdots
\bar{x}_1	\bar{x}_2	\bar{x}_3	\bar{x}_4

Notation:

Group index: $i \in \{1, 2, 3, 4\}$

Sample index: $j \in N_i = \text{size of group } i$

Data point x_{ij} : i th group, j th observation

Number of groups: $I = 4$

Total mean: \bar{x}

Group mean: \bar{x}_i

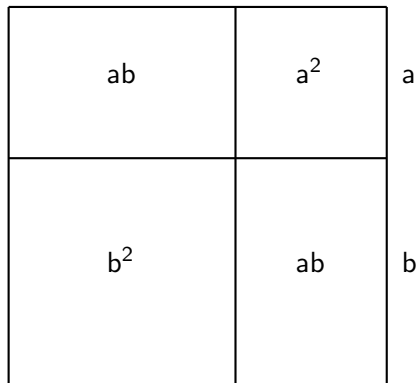
For any data point x_{ij} :

$$\begin{aligned}(x_{ij} - \bar{x}) &= (\bar{x}_i - \bar{x}) + (x_{ij} - \bar{x}_i) \\ \text{total residue} &= \text{group diff.} + \text{"error"}\end{aligned}$$

ANOVA question: does group membership influence the response variable?

Two variances

Reminder of high school algebra: $(a + b)^2 = a^2 + b^2 + 2ab$



Two variances

Data point x_{ij} :

$$(x_{ij} - \bar{x}) = (\bar{x}_i - \bar{x}) + (x_{ij} - \bar{x}_i)$$

Want sum of squared deviates for each group:

$$(x_{ij} - \bar{x})^2 = (\bar{x}_i - \bar{x})^2 + (x_{ij} - \bar{x}_i)^2 + 2(\bar{x}_i - \bar{x})(x_{ij} - \bar{x}_i)$$

Sum over elements in i th group:

$$\sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2 = \sum_{j=1}^{N_i} (\bar{x}_i - \bar{x})^2 + \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2 + \sum_{j=1}^{N_i} 2(\bar{x}_i - \bar{x})(x_{ij} - \bar{x}_i)$$

Two variances

Note that this term must be zero:

$$\sum_{j=1}^{N_i} 2(\bar{x}_i - \bar{x})(x_{ij} - \bar{x}_i)$$

Because:

$$(a) \quad \sum_{j=1}^{N_i} 2(\bar{x}_i - \bar{x})(x_{ij} - \bar{x}_i) = 2(\bar{x}_i - \bar{x}) \underbrace{\sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)}_0$$

$$(b) \quad \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i) = 0 \Leftrightarrow \bar{x}_i = \frac{\sum_{j=1}^{N_i} x_{ij}}{N_i}$$

Two variances

So we have:

$$\begin{aligned}\sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2 &= \sum_{j=1}^{N_i} (\bar{x}_i - \bar{x})^2 + \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2 \\ &\quad (+ \sum_{j=1}^{N_i} 2(\bar{x}_i - \bar{x})(x_{ij} - \bar{x}_i) = 0)\end{aligned}$$

Therefore:

$$\sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2 = \sum_{j=1}^{N_i} (\bar{x}_i - \bar{x})^2 + \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2$$

And finally we can sum over all groups:

$$\sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2 = \sum_{i=1}^I \sum_{j=1}^{N_i} (\bar{x}_i - \bar{x})^2 + \sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2$$

ANOVA terminology

$$\begin{array}{rclcl} (x_{ij} - \bar{x}) & = & (\bar{x}_i - \bar{x}) & + & (x_{ij} - \bar{x}_i) \\ \text{total residue} & = & \text{group diff.} & + & \text{"error"} \end{array}$$

$$\begin{array}{rclcl} \sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2 & = & \sum_{i=1}^I N_i (\bar{x}_i - \bar{x})^2 & + & \sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2 \\ \text{SST} & = & \text{SSG} & + & \text{SSE} \\ \text{Total Sum of Squares} & = & \text{Group Sum of Squares} & + & \text{Error Sum of Squares} \end{array}$$

$$\begin{array}{rclcl} (n - 1) & = & (I - 1) & + & (n - I) \\ \text{DFT} & = & \text{DFG} & + & \text{DFE} \\ \text{Total Degrees of Freedom} & = & \text{Group Degrees of Freedom} & + & \text{Error Degrees of Freedom} \end{array}$$

Variances are mean squared differences to the mean

Note that

SST/DFT: $\frac{\sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x})^2}{n-1}$ is a variance, and likewise

SSG/DFG: labelled **MSG** (“Mean square between groups”), and

SSE/DFE: labelled **MSE** (“Mean square error” or sometimes
“Mean square within groups”)

In ANOVA, we compare MSG (variance between groups) and MSE (variance within groups), i.e. we measure

$$F = \frac{\text{MSG}}{\text{MSE}}$$

If this F -value is large, differences between groups overshadow differences within groups.

Two variances

1) Estimate the pooled variance of the population (MSE):

$$\text{MSE} = \frac{\text{SSE}}{\text{DFE}} = \frac{\sum_{i=1}^I \sum_{j=1}^{N_i} (x_{ij} - \bar{x}_i)^2}{n - I} \quad \underline{\text{equiv}} \quad \frac{\sum_{i=1}^I \text{DF}_i \cdot s_i^2}{\sum_{i=1}^I \text{DF}_i}$$

In our example (Nederlands for anderstaligen):

$$\begin{aligned} \frac{\sum_{i=1}^I \text{DF}_i \cdot s_i^2}{\sum_{i=1}^I \text{DF}_i} &= \frac{(N_1 - 1)s_1^2 + (N_3 - 1)s_2^2 + (N_3 - 1)s_3^2 + (N_4 - 1)s_4^2}{(N_1 - 1) + (N_3 - 1) + (N_3 - 1) + (N_4 - 1)} \\ &= \frac{9 \cdot 66.22 + 9 \cdot 43.66 + 9 \cdot 34.99 + 9 \cdot 47.57}{9 + 9 + 9 + 9} \\ &= \frac{595.98 + 392.94 + 314.91 + 428.13}{36} = 48.11 \end{aligned}$$

Estimates the variance in groups (using DF), aka **within-groups estimate** of variance

2) Estimate the **between-groups** variance of the population (MSG):

$$\text{MSG} = \frac{\text{SSG}}{\text{DFG}} = \frac{\sum_{i=1}^I N_i (\bar{x}_i - \bar{x})^2}{I-1}$$

In our example (Nederlands for anderstaligen):

We had 4 group means: 25.0, 21.9, 23.1, 21.3, grand mean: 22.8

$$\text{MSG} = \frac{10 \cdot ((25-22.8)^2 + (21.9-22.8)^2 + (23.1-22.8)^2 + (21.3-22.8)^2)}{4-1} = 26.6$$

The **between-groups** variance (MSG) is an aggregate estimate of the degree to which the four sample means differ from one another

Interpreting estimates with F -scores

If H_0 is true, then we have two variances:

- ▶ Between-groups estimate: $s_{bg}^2 = 26.62$ and
- ▶ Within-groups estimate: $s_{wg}^2 = 48.11$

and their ratio $\frac{s_{bg}^2}{s_{wg}^2}$ follows an F -distribution with:

(# groups - 1) = 3 degrees of freedom for s_{bg}^2 and

(# observations - # groups) = 36 degrees of freedom for s_{wg}^2

In our example: $F(3, 36) = \frac{26.62}{48.11} = 0.55$

$P(F(3, 40) > 2.84) = 0.05$ (see tables), so there is no evidence of non-uniform behavior

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Variable NL_NIVO toets nl. voor anderstalige
By Variable GROUP gebied van afkomst

Analysis of Variance

Source	D.F.	Sum of Squares	Mean Squares	F Ratio	F Prob.
Between Groups	3	79.9	26.6	.55	.65
Within Groups	36	1731.9	48.1		
Total	39	1811.8			

No evidence of non-uniform behavior

Other questions

ANOVA $H_0: \mu_1 = \mu_2 = \dots = \mu_n$

But sometimes particular **contrasts** are important—e.g., are Europeans better (in learning Dutch)?

Distinguish (in reporting results):

- ▶ **prior** contrasts
questions asked before data is collected and analyzed
- ▶ **post hoc** (posterior) questions
questions asked **after** data collection and analysis
“data-snooping” is exploratory, cannot contribute to hypothesis testing

Questions asked **before** data collection and analysis—e.g., are Europeans better (in learning Dutch)?

Another way of putting this:

$$H_0 : \mu_{\text{Eur}} = \frac{1}{3}(\mu_{\text{Am}} + \mu_{\text{Afr}} + \mu_{\text{Asia}})$$
$$H_a : \mu_{\text{Eur}} \neq \frac{1}{3}(\mu_{\text{Am}} + \mu_{\text{Afr}} + \mu_{\text{Asia}})$$

Reformulation (SPSS requires this):

$$H_0 : 0 = -\mu_{\text{Eur}} + 0.33\mu_{\text{Am}} + 0.33\mu_{\text{Afr}} + 0.33\mu_{\text{Asia}}$$

Post-hoc questions

Assume H_0 is rejected: which means are distinct?

Data-snooping problem: in a large set, **some** distinctions are **likely** to be statistically significant

But we can still look (we just cannot claim to have **tested** the hypothesis)

We are asking whether $m_i - m_j$ is significantly larger, we apply a variant of the t -test

The relevant sd is $\sqrt{\frac{\text{MSE}}{n}}$ (differences among scores), but there is a correction since we're looking at a proportion of the scores in any one comparison

SD in post-hoc ANOVA questions

Standard deviation (among differences in groups i and j):

$$\text{sd} = \sqrt{\text{MSE} \times \frac{N_i + N_j}{N}} = \sqrt{48.1 \times \frac{10+10}{40}} = 4.9$$

$$t = \frac{\bar{x}_i - \bar{x}_j}{\text{sd} \cdot \sqrt{\frac{1}{N_i} + \frac{1}{N_j}}}$$

The critical t -value is calculated as $\frac{p}{c}$ where p is the desired significance level and c is the number of comparisons.

For pairwise comparisons: $c = \binom{I}{2}$

Post-hoc questions in SPSS

SPSS post-hoc 'Bonferroni'-searches among **all** groupings for statistically significant ones

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Variable  NL_NIVO      toets nl. voor anderstalige
By Variable  GROUP      gebied van afkomst
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Multiple Range Tests: Modified LSD (Bonferroni) test w. signif. level .05

The difference between two means is significant if

$MEAN(J) - MEAN(I) \geq 4.9045 * RANGE * \sqrt{1/N(I) + 1/N(J)}$

with the following value(s) for RANGE: 3.95

- No two groups significantly different at .05 level

Homogeneous Subsets (highest \& lowest means not sig. diff.)

Group	Azie	America	Africa	Europa
Mean	21.3	21.9	23.1	25.0

But in this case there are none (of course)

How to win at ANOVA

Note the ways in which the F -ratio increases (i.e., becomes more significant):

$$F = \frac{MSG}{MSE}$$

1. MSG increases: differences in means between groups grow larger
2. MSE decreases: overall variation within groups grows smaller

Two models for grouped data

$$x_{ij} = \mu + \epsilon_{ij}$$

$$x_{ij} = \mu + \alpha_i + \epsilon_{ij}$$

First model:

- ▶ no group effect
- ▶ each data point represents error (ϵ) around a mean (μ)

Second model:

- ▶ real group effect
- ▶ each data point represents error (ϵ) around an overall mean (μ), combined with a group adjustment (α_i)

ANOVA asks: is there sufficient evidence for α_i ?

Fall backs?

Suppose some cells are non-normal, or some standard deviations too large

Fall backs?

Suppose some cells are non-normal, or some standard deviations too large

- ▶ Kruskal-Wallis, non-parametric comparison of > 2 medians;
- ▶ apply (monotonic) transformation to reduce SD, perhaps improve fit to normality;
- ▶ trim most extreme 1% (or 5%) of data

Always report transformations or “trimming”!

Next week

Next week: factorial ANOVA