LEXICAL DIFFERENCES BETWEEN TUSCAN DIALECTS AND STANDARD ITALIAN: A SOCIOLINGUISTIC ANALYSIS USING GENERALIZED ADDITIVE MIXED MODELING

Martijn Wieling^a, Simonetta Montemagni^b, John Nerbonne^c and R. Harald Baayen^{a,d}

^aDepartment of General Linguistics, University of Tübingen, Germany, ^bIstituto di Linguistica Computationale "Antonio Zampolli", CNR, Italy, ^cDepartment of Humanities Computing, University of Groningen, the Netherlands, and ^dDepartment of Linguistics, University of Alberta, Canada wieling@gmail.com, simonetta.montemagni@ilc.cnr.it, j.nerbonne@rug.nl, harald.baayen@unituebingen.de

Abstract

This study uses a generalized additive mixed-effects regression model to predict lexical differences in Tuscan dialects with respect to standard Italian. We used lexical information for 170 concepts used by 2060 speakers in 213 locations in Tuscany. In our model, geographical position was found to be an important predictor, with locations more distant from Florence having lexical forms more likely to differ from standard Italian. In addition, the geographical pattern varied significantly for low versus high frequency concepts and older versus younger speakers. Younger speakers generally used variants more likely to match the standard language. Several other factors emerged as significant. Male speakers as well as farmers were more likely to use a lexical form different from standard Italian. In contrast, higher educated speakers used lexical forms more likely to match the standard. The model also indicates that lexical variants used in smaller communities are more likely to differ from standard Italian. The impact of community size, however, varied from concept to concept. For a majority of concepts, lexical variants used in smaller communities are more likely to differ from the standard Italian form. For a minority of concepts, however, lexical variants used in larger communities are more likely to differ from standard Italian. Similarly, the effect of the other community- and speaker-related predictors varied per concept. These results clearly show that the model succeeds in teasing apart different forces influencing the dialect landscape and helps us to shed light on the complex interaction between the standard Italian language and the Tuscan dialectal varieties. In addition, this study illustrates the potential of generalized additive mixed-effects regression modeling applied to dialect data.

Key words

Tuscan dialects, Lexical variation, Generalized additive modeling, Mixed-effects regression modeling

1. Introduction

In spite of their different origin and history, it is nowadays a widely acknowledged fact that traditional dialectology (to be understood here as dialect geography) and sociolinguistics (or urban dialectology) can be seen as two streams of a unique and coherent discipline: modern dialectology (Chambers and Trudgill, 1998). Chambers and Trudgill (1998:187-188) describe the convergence of these two historically separated disciplines as follows:

For all their differences, dialectology and sociolinguistics converge at their deepest point. Both are dialectologies, so to speak. They share their essential subject matter. Both fix the attention on language in communities. Prototypically, one has been centrally concerned with rural communities and the other with urban centres, but these are accidental differences, not essential ones and certainly not axiomatic. [...] A decade or two ago, it might have been possible to think that the common subject matter of dialectology and sociolinguistics counted for next to nothing. Now we know it counts for everything.

In practice, however, dialectology and sociolinguistics remain separate fields when considering the methods and techniques used for analyzing language variation and change.

Sociolinguistics - whose basic goal consists of identifying the social factors underlying the use of different variants of linguistic variables - adopted a quantitative approach to data analysis since its inception (e.g., Labov, 1966). Over time, different methods for the analysis of linguistic variation were developed, capable of modeling the joint effect of an increasing number of factors related to the social background of speakers (including age, gender, socio-economic status, etc.) and linguistic features. While early studies focused on simple relationships between the value of a linguistic variable and the value of a social variable (see e.g. Labov, 1966, 1972), over time more advanced statistical methods for the analysis of linguistic variation were developed. Since the 1970s, the most common method in sociolinguistic research has been logistic regression (Cedergren and Sankoff, 1974) and more recently, mixed-

effects regression models have been applied to socio-linguistic data (Johnson, 2009; Tagliamonte and Baayen, 2012, Wieling et al., 2011).

Traditional dialectology shows a different pattern. Since its origin in the second half of the 19th century, it typically relied on the subjective analysis of categorical maps charting the distribution of the different variants of a linguistic variable across a region. Only later, i.e. during the last forty years, quantitative methods have been applied to the analysis of dialect variation. This quantitative approach to the study of dialects is known as dialectometry (Séguy, 1973; Goebl, 1984, 2006; Nerbonne et al., 1996; Nerbonne, 2003; Nerbonne and Kleiweg, 2007). Dialectometric methods focus mostly on identifying the most important dialectal groups (i.e. in terms of geography) using an aggregate analysis of the linguistic data. The aggregate analysis is based on computing the distance (or similarity) between every pair of locations in the dataset based on the complete set of linguistic variables and by analyzing the resulting linguistic distance (or similarity) matrix using multivariate statistics to identify aggregate geographical patterns of linguistic variation.

While viewing dialect differences at an aggregate level arguably provides a more comprehensive and objective view than the analysis of a small number of subjectively selected features (Nerbonne, 2009), the aggregate approach has never fully convinced linguists of its use as it fails to identify the linguistic basis of the identified groups (see e.g. Loporcaro, 2009). By initially aggregating the values of numerous linguistic variables, traditional dialectometric analyses offer no direct method for testing whether and to what extent an individual linguistic variable contributes to observed patterns of variation. Recent developments in dialectometric research tried to reduce the gap between models of linguistic variation based on quantitative analyses and more traditional analyses based on specific linguistic features. Wieling and Nerbonne (2010, 2011) proposed a new dialectometric method, the spectral partitioning of bipartite graphs, to cluster linguistic varieties and simultaneously determine the underlying linguistic basis. This method, originally applied to Dutch dialects, was also successfully tested on English (Wieling et al., 2013) and Tuscan (Montemagni et al., 2012) dialects. Unfortunately, these methods still disregard social factors, and only take into account the influence of geography.

While some attempts have been made, social and spatial analyses of language are still far from being integrated. Britain (2002) reports that sociolinguistics fails to incorporate the notion of spatiality in its research. On the other hand, dialectometry mainly focuses on dialect geography and generally disregards social factors. The few exceptions indeed "prove" the proverbial rule. Montemagni et al. (2013) and Valls et al. (2013) included in their dialectometric analyses social factors concerning the difference between age classes or urban versus rural communities. Unfortunately, the effect of these social factors was evaluated by simply comparing maps visually, as opposed to statistically testing the differences. Another relevant aspect on which the sociolinguistic and dialectometric perspectives do not coincide concerns the role of individual features, which are central in sociolinguistics, but are typically and programmatically disregarded in dialectometry. These issues demonstrate that there is an increasing need for statistical methods capable of accounting for both the geographic and socio-demographic variation, as well as for the impact and role of individual linguistic features.

The present study is methodologically ambitious for its attempt to combine dialectometric and sociolinguistic perspectives along the lines depicted above. The statistical analysis methods we employ enable the incorporation of candidate explanatory variables based on social, geographical as well as linguistic factors, making it a good technique to facilitate the intellectual merger of dialectology and sociolinguistics (Wieling, 2012). The starting point is the study by Wieling, Nerbonne and Baayen (2011) who proposed a novel method using a generalized additive model in combination with a mixed-effects regression approach to simultaneously account for the effects of geographical, social and linguistic variables. They used a basic generalized additive model to represent the global geographical pattern, which was used in a second step as a predictor in their linear mixed-effects regression model. Their model predicted word pronunciation distances from the standard language to 424 Dutch dialects and it turned out that both the geographical location of the communities, as well as several location-related predictors (i.e. community size and average community age) and word-related factors (i.e. word frequency and category) were significant predictors. While the study of Wieling et al. (2011) includes social, lexical and geographical information, a drawback of their study is that they only considered a single speaker per location, limiting the potential influence of speakerrelated variables.

In this paper, we present an extended analytical framework which was tested on an interesting case study: Tuscan lexical variation with respect to standard Italian. There are three clear and important differences with respect to the study of Wieling et al. (2011). First, since the software available for generalized additive mixed-effects regression modeling has improved significantly since the study of Wieling et al. (2011), we are able to advance on their approach by constructing a single generalized additive mixed-effects regression model. This is especially beneficial as we are now in a position to better assess the effect of concept frequency, a variable which has largely been ignored from dialectological studies, but is highly relevant as it "[...] may affect the rate at which new words arise and become adopted in populations of speakers" (Pagel et al., 2007). Second, in this study we focus on lexical variation rather than variation in pronunciation. We therefore do not try to predict dialect distances, but rather a binary value indicating whether the lexicalization of a concept is different (1) or equal (0) with respect to standard Italian. A benefit of this approach is that it is more in line with standard sociolinguistic practice, which also focuses on binary distinctions. Third, as we take into account multiple speakers per location, we are at an improved position to investigate the contribution of speaker-related variables such as age and gender.

The Tuscan dialect case study we use to investigate the potential of this new method (integrating social, geographical and lexical factors) is a challenging one. In Italy a complex relationship exists between the standard language and dialects due to the history of this language and the circumstances under which Italy achieved political unification in 1861, much later than in most European countries. In Tuscany, a region with a special status among Italian dialects, the situation is even more complex as standard Italian is based on Tuscan, and in particular on the Florentine variety, which achieved national and international prestige from the fourteenth century onwards as a literary language and only later (after the Italian Unification, and mainly in the twentieth century) as a spoken language. However, standard Italian has never been identical to genuine Tuscan and is perhaps best described as an "abstraction" increasingly used for general communication purposes. The aim of this study,

therefore, is to investigate this particular relationship between Italian and Tuscan dialects. We focus on lexical variation in Tuscan dialects compared to standard Italian with the goal of defining the impact, role and interaction of a wide range of factors (i.e. social, lexical and geographical) in determining lexical choice by Tuscan dialect speakers. The study is based on a large set of dialect data, i.e. the lexicalizations of 170 concepts attested by 2060 speakers in 213 Tuscan varieties drawn from the corpus of dialectal data *Atlante Lessicale Toscano* ('Lexical Atlas of Tuscany', henceforth ALT; Giacomelli et al., 2000) in which lexical data have both a diatopic and diastratic characterization.

After discussing the special relationship between standard Italian and the Tuscan dialects in the next section, we will describe the Tuscan dialect dataset, followed by a more in-depth explanation of the generalized additive modeling procedure, our results and the implications of our findings.

2. Tuscan dialects and standard Italian

2.1. The notions of dialect and standard language in the Italian context

As pointed out by Berruto (2005), Italy's *dialetti* do not correspond to the same entity as e.g., the English dialects. Following the Coserian distinction among primary, secondary and tertiary dialects (Coseriu, 1980), the Italian dialects are to be understood as primary dialects, i.e. dialects having their own autonomous linguistic system, whereas the English dialects represent tertiary dialects, i.e. varieties resulting from the social and/or geographical differentiation of the standard language. Italian dialects – or, more technically, Italo-Romance varieties – thus do not represent varieties of Italian but independent 'sister' languages arisen from local developments of Latin (Maiden, 1995).

A similar 'sisterhood' relationship also exists between the Italian language and Italo-Romance dialects, because Italian has its roots in one of the speech varieties that emerged from spoken Vulgar Latin (Maiden and Parry, 1997), namely that of Tuscany, and more precisely the variety of Tuscan spoken in Florence. The importance of the Florentine variety in Italy was mainly determined by the prestige of the Florentine culture, and in particular the establishment of Dante, Petrarch and Boccaccio, who wrote in Florentine, as the 'three crowns' (*tre corone*) of Italian literature. The fact that standard Italian originated from the Florentine dialect centuries ago changes the type of relationship between standard Italian and Tuscan dialects to a kind of 'parental' relationship instead of a 'sisterhood' relationship. Clearly, this complicates matters with respect to the relationship between the Tuscan dialects and the standard Italian language, and this is the topic of the present study.

Standard Italian is unique among modern European standard languages. Even though it originated in the fourteenth century, it was not consolidated as a spoken national language until the twentieth century. For centuries, Italian was a written literary language, acquired through literacy when one learned to read and write, and was therefore only known to a minority of (literate) people. During this period, people spoke only their local dialect. For a detailed account of the rise of standard Italian the interested reader is referred to e.g. Migliorini and Griffith (1984). The particular nature of Italian as a literary language, rather than a spoken language, was recognized since its origin and has been widely debated from different (i.e. socio-economic, political and cultural) perspectives under the general heading of *questione della lingua* or 'language question'.

At the time of the Italian political unification in 1861 only a very small percentage of the population was able to speak Italian, with estimates ranging from 2.5% (De Mauro, 1963) to 10% (Castellani, 1982). Only during the second half of the 20th century real native speakers of Italian started to appear, as Italian started to be used by Italians as a spoken language in everyday life. Mass media (newspapers, radio and TV), education, and the introduction of compulsory military service played a central role in the diffusion of the Italian language throughout the country. According to recent statistics by the Italian National Census (*Istituto Nazionale di Statistica*, ISTAT) reported by Lepschy (2002), 98% of the Italian population is able to use their national language. However, dialects and standard Italian continue to coexist. For example, ISTAT data show that at the end of the 20th century (1996) 50% of the population used (mainly or exclusively) standard Italian to communicate with friends and colleagues, while this percentage decreased to 34% when communication with relatives was taken into account. More recently, Dal Negro and Vietti (2011)

presented a quantitative analysis of the patterns of language choice in present-day Italy on the basis of a national survey carried out by ISTAT in 2006. At the national level, they reported that 45.5% exclusively used Italian in a family setting, whereas 32.5% of the people alternated between dialectal and Italian speech, and 16% exclusively spoke in dialect (with the remaining ones using another language).

The current sociolinguistic situation of Italy is characterized by the presence of regional varieties of Italian (e.g., Berruto, 1989, 2005; Cerutti, 2011). Following the tripartite Coserian classification of dialects, these can be seen as tertiary dialects, i.e. varieties of the standard language that are spoken in different geographical areas. They differ both from each other and from standard Italian at all levels (phonetic, prosodic, syntactic and lexical), and represent the Italian actually spoken in contemporary Italy. Common Italian speakers generally speak a regional variety of Italian, referred to as regional Italian. The consequence of this is that there are no real native speakers of standard Italian. Not even a Tuscan or Florentine native speaker could be considered a native speaker of standard Italian, as in Tuscan or Florentine Italian features exist (such as the well-known Tuscan *gorgia*) which are not part of the standard Italian norm.

This clearly raises the question of what we mean by the standard Italian language. Generally speaking, a standard language is a fuzzy notion. Following Ammon (2004), the standard variety of a language can be seen as having a core of undoubtedly standard forms while also having fuzzy boundaries resulting in a complex gradation between standard and non-standard. In Italy, a new standard variety "neo-standard Italian" (Berruto, 1987) is emerging as the result of a restandardization process, which allows for a certain amount of regional differentiation. For the specific concerns of this study, aimed at reconstructing the factors governing the lexical choices of Tuscan speakers between dialect and standard language, we will refer to the core of undoubtedly standard forms as standard Italian. This is the only way to avoid interferences with the regional Italian spoken in Tuscany.

2.2. Previous studies on the relationship between standard Italian and Tuscan dialects

The specific relationship linking standard Italian and Tuscan dialects has been investigated in numerous studies. Given the goal of our research, we will only discuss those studies which focus on the lexical level.

The historical link between the Tuscan dialects and the standard Italian language causes frequent overlap between dialectal and standard lexical forms in Tuscany, and less frequent overlap in other Italian regions (Giacomelli, 1978). However, since Tuscan dialects have developed (for several centuries) along their own lines and independently of the (literary) standard Italian language, their vocabulary does not always coincide with that of standard Italian. Following Giacomelli (1975), the types of mismatch between standard Italian and the dialectal forms can be partitioned into three groups. The first group consists of Tuscan words which are used in literature throughout Italy, but are not part of the standard language (i.e. these terms usually appear in Italian dictionaries marked as 'Tuscanisms'). The second group consists of Tuscan words which *were* part of old Italian and are also attested in the literature throughout Italy, but have fallen into disuse as they are considered old-fashioned (i.e. these terms may appear in Italian dictionaries marked as 'archaisms'). The final group consists of Tuscan dialectal words which have no literary tradition and are not understood outside of Tuscany.

Here our goal is to investigate the complex relationship between standard Italian and the Tuscan dialects from which it originated on the basis of the data collected through fieldwork for the *Atlante Lessicale Toscano* (ALT). Previous studies have already explored the ALT dataset by investigating the relationship between Tuscan and Italian from the lexical point of view. Giacomelli and Poggi Salani (1984) based their analysis on the dialect data available at that time. Montemagni (2008), more recently, applied dialectometric techniques to the whole ALT dialectal corpus to investigate the relationship between Tuscan and Italian. In both cases it turned out that the Tuscan dialects overlap most closely with standard Italian in the area around Florence, expanding in different directions and in particular towards the southwest. Obviously, this observed synchronic pattern of lexical variation has the well-known diachronic explanation that the standard Italian language originated from the Florentine variety of Tuscan.

Montemagni (2008) also found that the observed patterns varied depending on the speaker's age: only 37 percent of the dialectal answers of the old speakers (i.e. born in 1920 or before) overlapped with standard Italian, while this percentage increased to 44 for the young speakers (i.e. born after 1945, when standard Italian started being progressively used). In addition, words having a larger geographical coverage (i.e. not specific to a small region), were more likely to coincide with the standard language than words attested in smaller areas. These first, basic results illustrate the potential of the ALT dataset (which we use here as well, and will be discussed in more detail in Section 3.1) to shed light on the complex relationship between standard Italian and Tuscan dialects.

3. Material

3.1. Lexical data

The lexical data used in this study was taken from the Atlante Lessicale Toscano (ALT), a specially designed regional atlas in which the dialectal data have a diatopic (geographic), diastratic (social) and diachronic characterization. The diachronic characterization covers only a few generations whose year of birth ranges from the end of the 19th century to the second half of the 20th century. It is interesting to note that only the younger ALT informants were born in the period when standard Italian started being used as a spoken language. ALT interviews were carried out between 1974 and 1986 in 224 localities of Tuscany. The localities were hierarchically organized according to their size, ranging from medium- sized urban centers (with the exclusion of big cities) to small villages and rural areas. In total there were 50 to 60 micro-areas, each placed around an urban center (for more details see Giannelli, 1978). In contrast to traditional atlases (typically relying on elderly and uneducated informants), the ALT includes 2193 informants which were selected with respect to a number of parameters ranging from age and socio-economic status to education and culture in order to be representative of the population of each location. The sample size for the individual localities ranges between 4 and 29 informants, depending on the population size. The temporal window covered by ALT makes this dataset particularly suitable to explore the complex relationship linking Tuscan dialects to standard Italian along several dimensions, i.e. across space, time and socially defined groups. The interviews were conducted by a group of trained fieldworkers who employed a questionnaire of 745 target items, designed to elicit variation mainly in vocabulary and semantics.

Since the compilation of the ALT questionnaire was aimed at capturing the specificity of Tuscan dialects and their relationships, concepts whose lexicalizations were identical to Italian (almost) everywhere in Tuscany were programmatically excluded (Giacomelli, 1978; Poggi Salani, 1978). This makes the ALT dataset particularly useful for better understanding the complex relationship linking the standard language and local dialects in the case the two did not coincide.

In this study, we focus on Tuscan dialects only, spoken in 213 out of the 224 investigated locations (see Figure 1; Gallo-Italian dialects spoken in Lunigiana and in small areas of the Apennines were excluded) reducing the number of informants to 2060. We used the normalized lexical answers to a subset of the ALT onomasiological questions (i.e. those looking for the attested lexicalizations of a given concept). Normalization was meant to abstract away from phonetic variation and in particular from productive phonetic processes, without removing morphological variation or variation caused by unproductive phonetic processes. Out of 460 onomasiological questions, we selected only those which prompted 50 or fewer distinct normalized lexical answers (the maximum in all onomasiological questions was 421 unique lexical answers). We used this threshold to exclude questions having many hapaxes as answers which did not appear to be lexical (a similar approach was taken by Montemagni, 2007). For example, the question looking for denominations of 'stupid' included 372 different normalized answers, 122 of which are hapaxes. These either represent productive figurative usages (e.g., metaphors such as cetriolo 'cucumber' and carciofo 'artichoke'), productive derivational processes (e.g., scemaccio and scemalone from the lexical root scemo 'stupid'), or multi-word expressions (e.g., *mezzo scemo* 'half stupid', *puro locco* 'pure stupid' and similar). From the resulting 195-item subset, we excluded a single adjective and twelve verbs (as the remaining concepts were nouns) and all twelve multi-word concepts. Our final

subset, therefore, consisted of 170 concepts and is listed in Table 1. Although ALT answers also include passive vocabulary, for this study we focused on lexical answers corresponding to active vocabulary only.

The normalized lexical forms in the ALT dataset still contained some morphological variation. In order to assess the pure lexical variation we abstracted away from variation originating in, e.g., assimilation, dissimilation, or other phonological differences (e.g., the dialectal variants *camomilla* and *capomilla*, meaning 'chamomile', have been treated as instantiations of the same normalized form), as well as from both inflectional and derivational morphological variation (e.g., inflectional variants such as singular and plural are grouped together). We compare these more abstract forms to the Italian standard. Note, however, that the effect of removing the morphological variation was relatively limited, as the results using the unaltered ALT normalized lexical forms were almost identical to the results based on the lexical forms where morphological variation was filtered out.



Figure 1. Geographical distribution of the 213 locations investigated in this study. The 'F', 'S' and 'P' mark the approximate locations of Florence, Siena and Pisa, respectively.

The list of standard Italian words denoting the 170 concepts was extracted from the online ALT dialectal resource (ALT-Web; available at http://serverdbt.ilc.cnr.it/altweb), where it had been created for query purposes, i.e. as a way for the user to identify the ALT question(s) corresponding to his or her research interests (see Cucurullo et al., 2006). This list, originally compiled on the basis of lexicographic evidence, was carefully reviewed by members of the Accademia della *Crusca*, the leading institution in the field of research on the Italian language in both Italy and the world, in order to make sure that it contained undoubtedly standard Italian forms and not old-fashioned or literary words originating in Tuscan dialects (see Section 2.2).

In every location multiple speakers were interviewed (see above) and therefore each normalized answer is anchored to a given location, but also to a specific speaker. As some speakers provided multiple distinct answers to denote a single concept, the total number of cases (i.e. concept-speaker-answer combinations) was 384,454.

As Wieling et al. (2011) reported a significant effect of word frequency on dialect distances from standard Dutch pronunciations (with more frequent words having a higher distance from standard Dutch, which was interpreted as a higher resistance to standardization), we obtained the concept frequencies (of the standard Italian lexical form) by extracting the corresponding frequencies from a large corpus of 8.4 million Italian unigrams (Brants and Franz, 2009). The corpus-based frequency ranking of these concepts was then compared to the Grande dizionario italiano dell'uso ('Comprehensive Dictionary of Italian Usage', GRADIT; De Mauro, 2000) which represents a standard usage-based reference resource for the Italian language including quantitative information on vocabulary use. In particular, a list of about 7000 high frequency words/concepts highly familiar to native speakers of Italian was identified in this dictionary, representing the so-called Basic Italian Vocabulary (BIV). It turned out that 59.4% of the concepts used in our study belonged to the BIV, whereas the remaining concepts refer to an old-fashioned and traditional world (19.4%), denote less common plants and animals (14.7%) or refer to kitchen tools (2.4%). The remaining 4.1% of the concepts represent a miscellaneous class. It is interesting to note that the classification of concepts with respect to this reference dictionary and the frequency data obtained from the large web corpus are aligned,

with our most frequent concepts being in the BIV and the low frequent concepts typically corresponding to old-fashioned and traditional notions as well as less common plants and animals.

3.2. Sociolinguistic data

The speaker information we obtained consisted of the speaker's year of birth, the gender of the speaker, the education level of the speaker (ranging from 1: illiterate or semi-literate to 6: university degree; for this variable about 1.3% of the values were missing) and the employment history of the speaker (in nine categories: farmer; craftsman; trader or businessman; executive or auxiliary worker; knowledge worker, manager or nurse; teacher or freelance worker; common laborer or apprentice; skilled or qualified worker; or non-professional status such as student, housewife or retired). Furthermore, we obtained the year of recording for every location and we extracted demographic information about each of the 213 locations from a website with statistical information about Italian municipalities (Comuni Italiani, 2011). We extracted the number of inhabitants (in 1971 or 1981, whichever year was closer to the year when the interviews for that location were conducted), the average income (in 2005; which was the oldest information available), and the average age (in 2007; again the oldest information available) in every location. While the information about the average income and average age was relatively recent and may not precisely reflect the situation at the time when the dataset was constructed (between 1974 and 1986), the global pattern will probably be relatively similar.

| abata | fir | cipresso | CUPPAGE | majalino | niglet | ramaiolo | ladla |
|--------------------|---------------------------|------------|--------------------|-------------|-----------------------|---------------------------|--|
| ubele | | cipresso | cypiess | 11 | pigiet | ramaioio | 12010 |
| acacia | acacia | cıspa | eye gum | mammella | breast | ramarro | green lizard |
| acino | grape | cocca | corner of tissue | mancia | tip | rana | frog |
| acquaio | sink | coperchio | cover | manciata | handful | ravanelli | radishes |
| albicocca | apricot | corbezzolo | arbutus | mandorla | almond | riccio | hedgehog |
| allodola | lark | corniolo | dogwood | mangiatoia | manger | <i>riccio</i> (castagna) | chestnut husk |
| alloro | laurel | crusca | bran | matassa | hank | ricotta | ricotta cheese |
| anatra | duck | cuneo | wedge | matterello | rolling pin | rosmarino | rosemary |
| angolo | ext. angle | dialetto | dialect | melone | melon | sagrato | churchyard |
| anguria | watermelon | ditale | thimble | mietitura | harvest | salice | willow |
| ane | hee | donnola | weasel | mirtillo | blueberry | saliva | saliva |
| arancia | orange | duna | dune | montone | ram | salsiccia | sallsage |
| urunciu | orange | auna | uulle | monione | Italian | saisiccia | sausage |
| aromi | aromas | edera | ivy | mortadella | sausage | scoiattolo | squirrel |
| aspide | asp | falegname | carpenter | Neve | snow | scorciatoia | shortcut |
| bigoncia | vat | faraona | guinea fowl | nocciola | hazelnut | scrofa | SOW |
| borraccina | moss | fiammifero | match | Oca | goose | seccatoio | squeegee |
| bottiglia | bottle | filare | spin | occhiali | glasses | sedano | celery |
| brace | embers | formica | ant | Orcio | jar | segale | rye |
| braciere | brazier | fragola | strawberry | orecchio | ear | sfoglia | pastry |
| braciola | chop | frangia | fringe | orzaiolo | stv | siero | serum |
| | enop | ji cingici | minge | 01201010 | | 51070 | Tuscan |
| bruco | caterpillar | frantoio | oil mill | Ovile | sheepfold | soprassata | salami made from the pig (offal) |
| cachi | khaki | fregatura | cheat | ovolo | royal agaric | spazzatura | garbage |
| caglio | rennet | fringuello | finch | padrino | godfather | spigolo | edge |
| calabrone | hornet | frinzello | badly done darn | pancetta | bacon | stollo | haystack pole |
| calderaio | tinker | fronte | front | pancia | belly | stoviglie | dishes |
| calvo | bald | fuliggine | soot | panzanella | Tuscan bread salad | straccivendolo | ragman |
| camomilla | chamomile | 90770 | magnie | papavero | poppy | susina | plum |
| cantina | cellar | galso | mulberry | paparere | robin | tacchino | turkey |
| caniina | centar | geiso | mulocity | pennosso | 100111 | lacenino | chopping |
| capezzolo | nipple | ghiandaia | jay | pigna | cone | tagliere | board |
| capocollo | cut from pork shoulder | ghiro | dormouse | pimpinella | pimpernel | talpa | mole |
| caprone | goat | ginepro | juniper | pinolo | pine seed | tartaruga | tortoise |
| carbonaio | charcoal | gomitolo | ball | pioppeto | poplar grove | trabiccolo (rotondo) | dome frame for bed |
| cascino | cheese mould | grandine | hail | pipistrello | bat | trabiccolo (allungato) | elongated frame for bed heating |
| castagnaccio | chestnut cake | grappolo | cluster | polenta | corn meal mush | trogolo | trough |
| castagneto | chestnut | grattugia | grater | pomeriggio | afternoon | truciolo | chip |
| cavalletta | grasshopper | grillo | cricket | presine | potholders | tuono | thunder |
| cetriolo | cucumber | idraulico | plumber | nrezzemolo | narslev | uncinetto | crochet |
| ciahatte | slinners | lampo | flash | Pula | chaff | ununa | hoopoe |
| | anppers | lanti | fra al-1 | 1 mu | flag | прири | hoon |
| ciccioli | greaves | ientiggini | ireckles | ршсе | <u> 11ea</u> | verro | DOAL |
| ciliegia | cherry | lucertola | lizard | pulcino | chick | vitalba | clematis |
| cimice | bug | lumaca | snail | puzzola | skunk | volpe | fox |
| <i>cintura</i> (f) | belt for woman | madrina | godmother | radice | root | | |
| cintura (m) | belt for man | maiale | pig | raganella | tree frog | | |

Table 1. List of all 170 lexical items included in this study including their English translation

4. Methods

4.1. Modeling the role of geography: generalized additive modeling

In contrast to a linear regression model in which a single predictor is linear in its effect on the dependent variable, in a generalized additive model (GAM) the assumption is relaxed so that the functional relation between a predictor and the response variable need not be linear. Instead, the GAM provides the user with a flexible toolkit for smoothing nonlinear relations in any number of dimensions. Consequently, the GAM is much more flexible than the simple linear regression model. In a GAM multiple predictors may be combined in a single smooth, yielding essentially a wiggly surface (when two independent variables are combined) or a wiggly hypersurface (when three or more independent variables are combined).

A GAM combines a standard linear model with regression coefficients β_0 , β_1 , ..., β_k with smooth functions s() for one or more predictors:

$$Y = \beta_0 + \beta_1 X_1 + \ldots + \beta_k X_k + s(X_i) + s(X_i, X_k) + \ldots$$

A suitable option to smooth a single predictor is to use cubic regression splines. These fit piecewise cubic polynomials (functions of the form $y = a + bx + cx^2 + dx^3$) to separate intervals of the predictor values. The transitions between the intervals (located at the knots) are ensured to be smooth as the first and second derivative are forced to be zero. The number of knots determines how smooth the curve is. Determining the appropriate amount of smoothing is part of the parameter estimation process.

To combine predictors which have the same scale (such as longitude and latitude), thin plate regression splines are a suitable choice. These fit a wiggly regression surface as a weighted sum of geometrically regular surfaces. When the predictors do not all have the same scale, tensor products can be used (Wood, 2006: 162). These define surfaces given marginal basis functions, one for each dimension of the smooth. The basis functions generally are cubic regression splines (but they can be thin plate regression splines as well) and the greater the number of knots for the different basis functions, the more wiggly the fitted regression surface will be. More information about the tensor product bases (which are implemented in the mgcv package for R) is provided by Wood (2006; Ch. 4). For the interested reader, the appendix shows the function call used to fit the complete generalized additive mixed-effects regression model. A more extended introduction about the use of generalized additive modeling in linguistics can be found in Baayen et al. (2010).

As it turns out, a thin plate regression spline is a highly suitable approach to model the influence of geography in dialectology, as geographically closer varieties tend to be linguistically more similar (e.g., see Nerbonne, 2010) and the dialectal landscape is generally quite smooth (note, however, that the method also can detect steep transitions between nearby geographical positions). Wieling et al. (2011) also used a generalized additive model to represent the global effect of geography, as this measure is more flexible than using e.g., distance from a certain point (Jaeger et al., 2011). In this study, we will take a more sophisticated approach, allowing the effect of geography to vary for concept frequency and speaker age. Furthermore, we will use a generalized additive *logistic* model, as our dependent variable is binary (in line with standard sociolinguistic practice using Varbrul; Cedergren and Sankoff, 1974). Logistic regression does not model the dependent variable directly, but it attempts to model the probability (in terms of logits) associated with the values of the dependent variable. A logit is the natural logarithm of the odds of observing a certain value (in our case, a lexical form different from standard Italian). Consequently, when interpreting the parameter estimates of our regression model, we should realize that these need to be interpreted with respect to the logit scale (i.e. the natural logarithm of the odds of observing a lexical form different from standard Italian). More detailed information about logistic regression is provided by Agresti (2007)

As an illustration of the GAM approach, Figure 2 presents the global effect of geography on lexical differences with respect to standard Italian. The complex wiggly surface shown here was modeled by a thin plate regression spline (Wood, 2003), which was also used by Wieling et al. (2011). The (solid) contour lines represent isolines connecting areas which have a similar likelihood of having a lexical form different from standard Italian. Note that the values here represent log-odds values (as we use logistic regression) and should be interpreted with respect to being different from standard Italian. This means that lower values indicate a smaller likelihood of

being different (intuitively it is therefore easiest to view these values as a distance measure from standard Italian). Consequently, the value -0.1 indicates that in those areas the lexical form is more likely to match the Italian standard (the probability is 0.475 that the lexical form is *different* from the Italian standard form) and the value 0.1 indicates the opposite (the probability is 0.525 that the lexical form is different from the Italian standard form). Correspondingly, darker shades of gray indicate a greater likelihood of having lexical forms identical to those in standard Italian, while lighter shades of gray represent a greater likelihood of having lexical forms different from those in standard Italian. We can clearly see that locations near Florence (indicated by the black star) tend to have lexical variants more likely to be identical to the standard Italian form. This makes sense as Italian originated from the Tuscan dialect spoken in Florence. The 27.49 estimated degrees of freedom invested in this general thin plate regression spline were supported by a Chi-square value of 1581 (p < 0.001).



Figure 2. Contour plot for the regression surface of predicting lexical differences from standard Italian as a function of longitude and latitude obtained with a generalized additive model using a thin plate regression spline. The (black) contour lines represent isolines, darker shades of gray (lower values) indicate a smaller likelihood of having a lexical form different from standard Italian, while lighter shades of gray (higher values) represent locations with a greater likelihood of having a lexical form different from standard Italian. The black star marks the location of Florence. The white squares indicate combinations of longitude and latitude for which there is no (nearby) data.

As Wieling et al. (2011) found that the effect of word frequency on (Dutch) dialect distances varied per location, we initially created a three-dimensional smooth (longitude x latitude x concept frequency), allowing us to assess the concept frequency-specific geographical pattern of lexical variation with respect to standard Italian. For example, it might be that the geographical pattern presented in Figure 2, may hold for concepts having an average frequency, but might be somewhat different for concepts with a low as opposed to a high frequency. As our initial analyses revealed that this pattern varied depending on speaker age, we also included the speaker's year of birth in the smooth, resulting in a four-dimensional smooth (longitude x latitude x concept frequency x speaker's year of birth). We model this four-dimensional smooth by a tensor product. In the tensor product, we model both longitude and latitude with a thin plate regression spline (as this is suitable for combining isotropic predictors and also in line with the approach of Wieling et al., 2011), while the effect of concept frequency and speaker's year of birth are modeled by two separate cubic regression splines.

4.2. Mixed-effects modeling

A generalized additive *mixed-effects* regression model distinguishes between fixed and random-effect factors. Fixed-effect factors have a small number of levels exhausting all possible levels (e.g., gender is either male or female). Random-effect factors, in contrast, have levels sampled from a much large population of possible levels. In our study, concepts, speakers and locations are random-effect factors, as we could have included many other concepts, speakers or locations. By including random-effect factors, the model can take the systematic variation linked to these factors into account. For example, some concepts will be more likely to be different from standard Italian than others (regardless of location) and some locations (e.g., near Florence) or speakers will be more likely to use lexical variants similar to standard Italian (across all concepts). These adjustments to the population intercept (consequently identified as 'random intercepts') can be used to make the regression formula more precise for every individual location and concept. It is also possible that there is variability in the effect a certain predictor has. For example, while the general effect of community size might be negative (i.e. larger communities have lexical variants more likely to match the standard Italian form), there may be significant variability for the individual concepts. While most concepts will follow the general pattern, some concepts could even exhibit the opposite pattern (i.e. being more likely to match the standard Italian form in smaller communities). In combination with the by-concept random intercepts, these by-concept random slopes make the regression formula for every individual concept as precise as possible. Furthermore, taking this variability into account prevents type-I errors in assessing the significance of the predictors of interest. The significance of random-effect factors in the model was assessed by the Wald test. More information and an introduction to mixed-effects regression models is provided by Baayen et al. (2008).

In our analyses, we considered the three aforementioned random-effect factors (i.e. location, speaker and concept) as well as several other predictors besides the (concept frequency and speaker age-specific) geographical variation. The additional speaker-related variables we included were gender, education level and employment history (coded in 9 binary variables denoting if a speaker has had each specific job or not). The demographic variables we investigated were community size, average community age, average community income, and the year of recording.

To reduce the potentially harmful effect of outliers, several numerical predictors were log-transformed (i.e. community size, average age, average income, education level and concept frequency). We scaled all numerical predictors by subtracting the mean and dividing by the standard deviation in order to facilitate the interpretation of the fitted parameters of the statistical model.

4.3. Combining mixed-effects regression and generalized additive modeling

In contrast to the approach of Wieling et al. (2011), where they first created a separate generalized additive model (similar to the one illustrated in Figure 2) and used the fitted values of this model as a predictor in a mixed-effects regression model, we are able to create a single generalized additive mixed-effects regression model, which estimates all parameters simultaneously. As the software to construct a generalized

additive model is continuously evolving, this approach was not possible previously. The specification of our generalized additive mixed-effects regression model using the mgcv package for R is shown in the appendix.

5. Results

We fitted a generalized additive mixed-effects logistic regression model, step by step removing predictors that did not contribute significantly to the model. In the following we will discuss the specification of the model including all significant predictors and verified random-effect factors.

Our response variable was binary with a value of 1 indicating that the lexical form was different from the standard Italian form and a value of 0 indicating that the lexical form was equal to standard Italian. The coefficients and the associated statistics of the significant fixed-effect factors and linear covariates are presented in Table 2. To allow a fair comparison of the effects of all predictors, we included a measure of effect size by specifying the increase or decrease of the likelihood of having a non-standard Italian lexical form (in terms of logits) when the predictor increased from its minimum to its maximum value. Table 3 presents the significance of the four-dimensional smooth term (modeling the concept frequency and speaker age-dependent geographical pattern) and Table 4 lists the significant random-effects structure of our model.¹

To evaluate the goodness of fit of the final model (see Tables 2 to 4), we used the index of concordance *C*. This index is also known as the receiver operating characteristic curve area 'C' (see, e.g., Harrell, 2001). Values of *C* exceeding 0.8 are generally regarded as indicative of a successful classifier. According to this measure, the model performed well with C = 0.82.

¹ The effect of removing the morphological variation (see Section 3.1) was relatively limited, as the results on the basis of the original data were mainly identical to the results shown in Tables 3, 4 and 5 (where morphological variation was removed). The only difference was that in the dataset including the morphological variation speakers who had a teaching or freelance profession were significantly (p < 0.05) more likely to use a standard Italian form than those who had another profession. This variable was not significant (p = 0.12) in the further normalized data (i.e. abstracting away from morphological as well as formal variation).

| | Estimate | Std. error | z-value | <i>p</i> -value | Eff. size |
|-----------------------|----------|------------|---------|-----------------|-----------|
| Intercept | -0.4184 | 0.1266 | -3.306 | < 0.001 | |
| Community size (log) | -0.0587 | 0.0224 | -2.621 | 0.009 | -0.3689 |
| Male gender | 0.0378 | 0.0128 | 2.948 | 0.003 | 0.0378 |
| Farmer profession | 0.0458 | 0.0169 | 2.711 | 0.007 | 0.0458 |
| Education level (log) | -0.0684 | 0.0126 | -5.433 | < 0.001 | -0.2762 |

Table 2. Significant parametric terms of the final model. A positive estimate indicates that a higher value for this predictor increases the likelihood of having a non-standard Italian lexical form, while a negative estimate indicates the opposite effect. Effect size indicates the increase or decrease of the likelihood of having a non-standard Italian lexical form when the predictor value increases from its minimum to its maximum value (i.e. the complete range).

| | Est. d.o.f. | Chi. sq. | <i>p</i> -value |
|---------------------------------|-------------|----------|-----------------|
| Geography x concept frequency x | 227.79 | 3283 | < 0.001 |
| speaker's year of birth | | | |

Table 3. Significant smooth term of the final model. The estimated degrees of freedom of the smooth term is indicated, as well as its significance in the model. Figure 3 shows the visualization.

| Factors | Random effects | Std. dev. | <i>p</i> -value |
|----------|-------------------------------------|-----------|-----------------|
| Speaker | Intercept | 0.0108 | 0.006 |
| Location | Intercept | 0.1879 | < 0.001 |
| Concept | Intercept | 1.5962 | < 0.001 |
| | Year of recording | 0.2838 | < 0.001 |
| | Community size (log) | 0.1787 | < 0.001 |
| | Average community income (log) | 0.2661 | < 0.001 |
| | Average community age (log) | 0.2414 | < 0.001 |
| | Farmer profession | 0.1027 | < 0.001 |
| | Executive or auxiliary worker prof. | 0.0654 | < 0.001 |
| | Education level (log) | 0.1256 | < 0.001 |
| | Male gender | 0.0802 | < 0.001 |

Table 4. Significant random-effect parameters of the final model. The standard deviation indicates the amount of variation for every random intercept and slope.

5.1. Geographical variation and lexical predictors

Inspecting Table 3, it is clear that geography is a very strong predictor, and it varied significantly with concept frequency and speaker age. We validated that the geographical pattern was necessary by comparing the AIC values (Akaike Information Criterion – the AIC indicates the relative goodness of fit of the model, with lower values signifying an improved model; Akaike, 1974). Including geography was necessary as the AIC for a model without geography (but including all predictors

and random-effect factors shown in Table 2 and Table 4) was, whereas the AIC for the model including a simple geographical smooth decreased (i.e. improved) to 392903. Furthermore, varying the geographical effect by speaker age further reduced the AIC to 392727, while varying it by word frequency resulted in an AIC of 390683. The best model (with an AIC of 390474) was obtained when the geographical effect varied depending on word frequency and speaker age. Figure 3 visualizes the geographical variation related to concept frequency and speaker age. Lighter shades of gray indicate a greater likelihood of having a lexical form different from standard Italian.

The three graphs to the left present the geographical patterns for the older speakers, while those to the right present the geographical patterns for the younger speakers. When going from the top to bottom, the graphs show the geographical pattern for increasing concept frequency.

The first observation is that all graphs show the same general trend according to which speakers from Florence (marked by the star) or the area immediately surrounding it are more likely to use a standard Italian form than the speakers from the more peripheral areas. This makes sense as standard Italian originated from Florence. Note, however, that the likelihood of using a standard Italian form varies significantly depending on the age of the speakers and the frequency of concepts.

With respect to the age of the speakers, comparing the left and right graphs yields a straightforward pattern: in all cases, the right graphs are much darker than the left ones, indicating that the younger speakers are much more likely to use a standard Italian form. Whereas the right graphs can thus be taken to reflect the standardizing effect deriving from the increased use of the standard Italian language, the left graphs are closer to the original pattern of Tuscan dialect variation characterized by a more limited influence of the standard language. This can be seen as following from the fact that the older ALT speakers were born between the end of the 19th century and the beginning of the 20th century. The two series of graphs can thus be seen as separate windows on different stages of the spreading of standard Italian in a region where the relationship between dialect and standard language is particularly complex due to the 'parental' relationship linking the two.

Let us now consider the effect of concept frequency. For the older speakers, we observe that the lexicalizations of high frequency concepts are less likely to be identical to standard Italian than those of low frequency concepts (i.e. the graph of the high frequency concepts is lighter than the graph of the low frequency concepts).

As the standard language also influences dialectal variation in older speakers, the left graphs in Figure 3 show that this is less effective for the higher frequency words. This is in line with previous research of Pagel et al. (2007), who found that words denoting frequently used concepts are less prone to be replaced (possibly because they are better entrenched in memory and therefore more resistant to lexical replacement). Furthermore, Wieling et al. (2011) also reported a resistance to standardization for high frequency words in Dutch dialects.

For the low frequency concepts, older speakers are more likely to use the standard Italian word. This, in our opinion, should not be seen as the result of standardization (at least in the majority of cases), but rather of the fact that low frequency concepts in our dataset typically belong to an obsolete, progressively disappearing rural world (some examples are *bigoncia* 'vat', *seccatoio* 'squeegee', and *stollo* 'haystack pole'). In this case the specific terms used in central Tuscan dialects to refer to these concepts are part of the standard Italian vocabulary.

The results recorded for older speakers suggest that the overlap between dialectal and standard lexical forms in Tuscany is not evenly distributed according to concept frequency. Overlap with standard Italian was most common for low frequency concepts, whereas high frequency concepts were more likely to be different from standard Italian. However, an in-depth explanation of the reasons underlying this state of affairs would require further analysis and data and therefore goes beyond the scope of this paper.



Figure 3. Contour plots for the regression surface of predicting lexical differences from standard Italian as a function of longitude, latitude, concept frequency, and speaker age obtained with a generalized additive model. The (black) contour lines represent isolines, darker shades of gray (lower values)

indicate a smaller lexical 'distance' from standard Italian (i.e. a smaller likelihood of having a lexical form different from standard Italian), while lighter shades of gray (higher values) represent locations with a larger lexical 'distance' from standard Italian. The star marks the location of Florence. The left plots visualize the results for older speakers (two standard deviations below the mean year of birth of 1931, i.e. 1888), while the right plots show those for the younger speakers (two standard deviations above the mean year of birth of 1931, i.e. 1975). The top row visualizes the contour plots for low frequency concepts (two standard deviations below the mean), the middle row for concepts having the mean frequency, and the bottom row for high frequency concepts (two standard deviations above the mean). The white squares in each graph indicate combinations of longitude and latitude for which there is no (nearby) data. See the text for interpretation.

For the younger speakers, a slightly different pattern can be observed. While the high frequency concepts are less likely to be identical to standard Italian than the mean frequency concepts (due to high frequency concepts being more resistant to change; Pagel et al., 2007), the low frequency concepts are also less likely to be identical to standard Italian. A possible explanation for this pattern is that, as previously stated, the low frequency concepts mostly consist of words from a disappearing rural world. Younger speakers might lack specific words for denoting these concepts and use more general terms instead (mismatching with the standard Italian form).

To conclude, we can state that the patterns of lexical choice between standard Italian and dialect by Tuscan speakers visually represented in Figure 3 do not purely show the effect of the standard Italian language on the Tuscan varieties, but also the complex diachronic relationship holding between the Florentine variety and the standard Italian language.

5.2. Speaker-related predictors

When inspecting Figure 3, it is clear that older speakers were much more likely to use forms different from standard Italian than younger speakers. This result is not unexpected as younger speakers tend to converge to standard Italian. In addition, we found clear support for the significance of gender. Men were much more likely to use non-standard forms than the standard Italian form. This finding is also not surprising, given that men generally use a higher frequency of non-standard forms than women (Cheshire, 2002). Analogous gender differences were reported for Tuscany by Cravens and Giannelli (1995) for what concerns the spread of intervocalic spirantization of /p/, /t/ and /k/ (i.e. the so-called Tuscan *gorgia*) as well as by Binazzi (1996) with respect to the use of dialectal words as opposed to standard Italian in Florence. Similarly, farmers were also found to be more likely to use non-standard forms. A reasonable explanation for this is that people living in rural areas (as farmers tend to do, given the nature of their work) generally favor non-standard forms and are less exposed to other language varieties (e.g., Chambers and Trudgill, 1998). The final significant speaker-related variable was education level. Higher educated speakers used forms more likely to be identical to the Italian standard. Again, this finding is not unforeseen as higher educated people tend to use more standard forms (e.g., Gorman, 2010).

As shown in Table 4, the effect of all speaker-related variables varied per word. For example, Figure 4 shows the effect of education level per word. Words such as *upupa* 'hoopoe' (a bird species) and *abete* 'fir' follow the general pattern (with higher educated speakers being more likely to use a standard form), while words such as *verro* 'boar' and *cocca* 'corner of a tissue' show the opposite behavior (with less educated people being more likely to use the standard form). As remarked before, taking these by-concept random slopes into account allows us to more reliably assess the general effect of the fixed-effect predictors.

5.3. Demographic predictors

Of all demographic predictors (i.e. the community size, the average community income and the average community age) only the first was a significant predictor in the general model. Larger communities were more likely to have a lexical variant identical to standard Italian (i.e. the estimate in Table 2 is negative). A possible explanation for this finding is that people tend to have weaker social ties in urban communities, which causes dialect leveling (Milroy, 2002), i.e. socially or locally marked variants tend to be leveled in favor of the standard language in conditions of social or geographical mobility and the resulting dialect contact.



Concepts sorted by the effect of education level

Figure 4. By-concept random slopes of education level. The concepts are sorted by the value of their education level coefficient (i.e. the effect of education level of the speakers). The strongly negative coefficients (bottom left) are associated with concepts that are more likely to be identical to standard Italian for higher educated speakers, while the positive coefficients (top right) are associated with concepts that are more likely to be different from standard Italian for higher educated speakers. The model estimate (see Table 2) is indicated by the dashed line.

The other demographic predictors, average age and average income, were not significant in the general model. In the study of Wieling et al. (2011) on Dutch dialects, average age was identified as a significant predictor of pronunciation distance from standard Dutch, while average income was not. The effect of average community age may be less powerful in our study, because we also included speaker age (which is much more suitable to detect the influence of age). In line with Wieling et al. (2011), the effect of average income pointed to a negative influence (with richer communities having lexical variants closer to the standard), but not significantly so (p = 0.5). Also note that year of recording was not significant as a fixed-effect predictor in the general model, which is likely caused by the relatively short time span (with respect to lexical change) in which the data was gathered.

All demographic variables (i.e. community size, average income and average age) as well as year of recording showed significant by-concept variation. For example, Figure 5, illustrating the effect of community size, shows some concepts (e.g., *trabiccolo* 'elongated frame for bed heating' and *mirtillo* 'blueberry') which are more likely to be identical to standard Italian in larger communities (i.e. consistent with the general pattern; the model estimate is indicated by the dashed line), while others behave in completely opposite fashion (i.e. *frinzello* 'badly done darn' and *nocciola* 'hazelnut') and are more likely to be different from standard Italian in larger communities.



Concepts sorted by the effect of community size

Figure 5. By-concept random slopes of community size. The concepts are sorted by the value of their community size coefficient (i.e. the effect of community size). The strongly negative coefficients (bottom left) are associated with concepts that are more likely to be identical to standard Italian in larger communities, while the positive coefficients (top right) are associated with concepts that are more likely to be different from standard Italian in larger communities. The model estimate (see Table 2) is indicated by the dashed line.

6. Discussion

In this study we have used a generalized additive model to identify the factors influencing the lexical choice of Tuscan speakers between dialect and standard Italian forms. We found clear support for the importance of speaker gender, speaker education level, speaker profession (i.e. being a farmer), community size, as well as geography, which varied significantly depending on concept frequency and speaker age. In addition, we illustrated that the mixed-effects regression approach enabled a detailed investigation of individual concepts. By simultaneously capturing the diatopic, diastratic and diachronic (though restricted to only a few generations) dimensions of variation and by permitting the analysis of individual linguistic features, we can claim that the proposed method is successful in combining the dialectometric and sociolinguistic perspectives in the analysis of dialectal lexical data.

The method was tested on a challenging case study focusing on the complex relationship between Tuscan dialects and standard Italian on the basis of the data gathered for the Atlante Lessicale Toscano, which turned out to offer an interesting and unique window into the complex interplay of diachronic and synchronic variation. The results which have emerged from our analysis of the ALT corpus shed new light on the typology, impact and role of a wide range of factors underlying the lexical choices by Tuscan speakers. Previous studies, based both on individual words (Giacomelli and Poggi Salani, 1984) and on aggregated data (Montemagni, 2008), provided a flat view according to which Tuscan dialects overlap most closely with standard Italian in the area around Florence, with expansions in different directions and in particular towards the southwest. Montemagni's (2008) aggregate analysis illustrated that a higher likelihood of using standard Italian was connected with speaker age and geographical coverage of words. In this study, however, a more finely articulated picture emerged. For example, we have shown that concept frequency also plays an important role, with more frequent concepts being more resistant to change.

On the demographic side, apart from observing that younger speakers were more likely to use forms identical to standard Italian, we found a significant effect of speaker gender, profession and education level (with male speakers, lower educated speakers, and farmers using lexical forms more likely to be different from standard Italian). Our gender-based findings thus provide further evidence supporting the Labovian Sex/Prestige pattern (Labov, 1966). In addition, we observed that larger communities are more likely to use standard Italian vocabulary than smaller communities.

Last but not least, because of the temporal window covered by the ALT dataset it was possible to keep track of the spreading of standard Italian and its increasing use as a spoken language. Real standardization effects could only be observed with respect to younger speakers, whereas older generations turned out to prefer dialectal variants, especially for high frequency concepts.

A limitation of this study is that it proceeded from dialect atlas data, which inherently suffers from a sampling bias. Furthermore, to keep the analysis tractable and focus on purely lexical variation we selected a subset of the data from the dialect atlas. While still having a relatively large number of items, our dataset only consisted of nouns. As the influence of word category might also vary geographically (see Wieling et al., 2011), further research is necessary to see if the results of this study extend to other word categories.

Another interesting line of research which might be worth pursuing would be to resort to a more sensitive distance measure with respect to standard Italian such as the Levenshtein (or edit) distance, rather than the binary lexical difference measure used in this study. In this case, lexical differences which are closely related (i.e. in the case of lexicalized analogical formations) can be distinguished from deeper lexical differences (e.g. due to a different etymon).

Acknowledgements

The research reported in this paper was carried out in the framework of the Short Term Mobility program of international exchanges funded by CNR (Italy).

References

- Agresti, A. (2007). An *introduction to categorical data analysis*. John Wiley & Sons, Hoboken, NJ, 2nd edition.
- Akaike H. (1974). A new look at the statistical identification model. *IEEE Transactions on Automatic Control*, 19, 716-723.
- Ammon, U. (2004). Standard variety. In U. Ammon, N. Dittmar, K.J. Mattheier and P. Trudgill (eds.), *Sociolinguistics. An International Handbook of the Science of Language and Society*, 2nd edn., vol. 1, Berlin and New York: Mouton de Gruyter, pp. 273–283.
- Baayen, R.H., D.J. Davidson and D.M. Bates (2008). Mixed-effects modeling with crossed random effects for subjects and items. *Journal of Memory and Language*, 59(4), 390-412.
- Baayen, R.H, Kuperman, V. and Bertram, R. (2010) Frequency Effects in Compound Processing. In Sergio Scalise and Irene Vogel (eds). *Compounding*. Benjamins, Amsterdam / Philadelphia, 257-270.
- Berruto, G. (1989). Main topics and findings in Italian sociolinguistics. International *Journal of the Sociology of Language*, 76 [Special issue: Italian sociolinguistics: Trends and issues], 5–30.
- Berruto, G. (2005). Dialect/standard convergence, mixing, and models of language contact: the case of Italy. In P. Auer, F. Hinskens and P. Kerswill (eds.), *Dialect change. Convergence and divergence in European languages.* Cambridge: Cambridge University Press, pp. 81–97.
- Binazzi, N. (1996). Giovani uomini e giovani donne di fronte al lessico della tradizione: risultati di un'analisi sul campo. In G. Marcato (ed.), Donna e linguaggio, Atti del Convegno Internazionale di Studi, Sappada/Plodn, 26-30 giugno 1995, Padova, Cleup, pp. 569-579.
- Brants, T. and A. Franz (2009). Web 1T 5-gram, 10 European Languages Version 1. Linguistic Data Consortium, Philadelphia.
- Britain, D. (2002). Space and spatial diffusion. In J. Chambers, P. Trudgill and N. Schilling-Estes (eds.), *The Handbook of Variation and Change*, Oxford: Blackwell, pp. 603-637.
- Castellani, A. (1982). Quanti erano gli italofoni nel 1861? *Studi Linguistici Italiani*, 8/1, Roma, Salerno Editrice, pp. 3-26.

- Cedergren, H.; and Sankoff, D. (1974). Variable rules: Performance as a statistical reflection of competence. *Language*, *50*, 333-355.
- Cerruti, M., (2011). Regional Varieties of Italian in the Linguistic Repertoire. International Journal of the Sociology of Language, 210, Walter de Gruyter, pp. 9-28.
- Comuni Italiani (2011). Informazioni e dati statistici sui comuni in Italia, le province e le regioni italiane. Sito ufficiale, CAP, numero abitanti, utili link. http://www.comuni-italiani.it/. Last accessed: 2011-05-23.
- Chambers, J.K. and P. Trudgill (1998). *Dialectology*. Second edition. Cambridge University Press, Cambridge.
- Cheshire, J. (2002). Sex and gender in variationist research. In J.K.Chambers, P. Trudgill and N. Schilling-Estes (eds.) Handbook of Language Variation and Change. Blackwell, Oxford, pp. 423-43.
- Coltheart, M. (1981). The MRC Psycholinguistic Database. *The Quarterly Journal of Experimental Psychology Section A*, 33(4), 497-505.
- Coseriu, E. (1980). "Historische Sprache" und "Dialekt". In J. Göschel, I. Pavle and K. Kehr (eds.), *Dialekt und Dialektologie*, Wiesbaden: Steiner, pp. 106–122.
- Cravens, T.D., L. Giannelli (1995). Relative salience of gender and class in a situation of multiple competing norms. *Language Variation and Change*, 7, 261-285.
- Cucurullo, S., S. Montemagni, M. Paoli, E. Picchi and E. Sassolini (2006). Dialectal resources on-line: the ALT-Web experience. *Proceedings of the 5th International Conference on Language Resources and Evaluation (LREC-2006)*, Genova, Italy, 24-26 May 2006, pp. 1846-1851.
- Dal Negro, S. and A. Vietti (2011). Italian and Italo-Romance dialects. *International Journal of the Sociology of Language*, 210, 71-92.
- De Mauro, T. (1963). Storia linguistica dell'Italia unita. Bari-Roma, Laterza.
- De Mauro, T. (2000). Grande dizionario italiano dell'uso. Torino, UTET.
- Giacomelli, G. (1975). Dialettologia toscana. *Archivio glottologico italiano*, 60, 179-191.
- Giacomelli, G. (1978). Come e perchè il questionario. In G. Giacomelli et al. (eds.), Atlante lessicale toscano - Note al questionario, Firenze, Facoltà di Lettere e Filosofia, 19-26.

- Giacomelli, G., L. Agostiniani, P. Bellucci, L. Giannelli, S. Montemagni, A. Nesi, M. Paoli, E. Picchi and T. Poggi Salani (2000). *Atlante Lessicale Toscano*. Lexis Progetti Editoriali, Roma.
- Giacomelli, G. and T. Poggi Salani (1984). Parole toscane. *Quaderni dell'Atlante Lessicale Toscano*, 2(3), 123-229.
- Giannelli, L. (1978). L'indagine come ricerca delle diversità. In G. Giacomelli et al. (eds.), Atlante lessicale toscano - Note al questionario, Firenze, Facoltà di Lettere e Filosofia, 35-50.
- Goebl, H., (1984). Dialektometrische Studien: Anhand italoromanischer, rätoromanischer und galloromanischer Sprachmaterialien aus AIS und ALF, Tübingen, M. Niemeyer.
- Goebl, H., (2006). Recent Advances in Salzburg Dialectometry. *Literary and Linguistic Computing*, 21(4), 411-435.
- Gorman, K. (2010). The consequences of multicollinearity among socioeconomic predictors of negative concord in Philadelphia. In Lerner, M. (ed.) University of Pennsylvania Working Papers in Linguistics, volume 16, issue 2, pp. 66–75.
- Harrell, F (2001). Regression modeling strategies. Springer, Berlin.
- Jaeger, T. F., P. Graff, B. Croft and D. Pontillo (2011). Mixed effect models for genetic and areal dependencies in linguistic typology: Commentary on Atkinson. *Linguistic Typology*, 15(2), 281-319
- Johnson, D.E. (2009). Getting off the GoldVarb Standard: Introducing Rbrul for Mixed-Effects Variable Rule Analysis. *Language and Linguistic Compass*, 3(1), 359–383.
- Labov, W. (1966). *The Social Stratification of English in New York City*. Center for Applied Linguistics, Washington, DC.
- Labov, W. (1972). *Sociolinguistic Patterns*. University of Pennsylvania Press, Philadelphia, PA.
- Lepschy, G. (2002). *Mother Tongues & other Reflections on the Italian Language*. University of Toronto Press, Toronto.
- Loporcaro, M. (2009). Profilo linguistico dei dialetti italiani, Roma-Bari, Laterza.
- Maiden, M. (1995). A Linguistic History of Italian. Longman, London.
- Maiden, M. and M. Parry (1997). The Dialects of Italy. Routledge, London.
- Migliorini, B. and T.G. Griffith (1984). *The Italian language*. Faber and Faber, London.

- Milroy, L. (2002). Social Networks. In J. Chambers, P. Trudgill and N. Schilling-Estes (eds.), *The Handbook of Language Variation and Change*. Blackwell Publishing Ltd., pp. 549-572.
- Montemagni, S., M. Wieling, B. de Jonge and J. Nerbonne (2012). Patterns of Language Variation and Underlying Linguistic Features: A New Dialectometric Approach. In Bianchi, P., De Blasi, N., De Caprio, C. and Montuori, F. (eds), La variazione nell'italiano e nella sua storia. Varietà e varianti linguistiche e testuali. Atti dell'XI Congresso SILFI (Società Internazionale di Linguistica e Filologia Italiana). Firenze: Franco Cesati Editore, vol. II, pp. 879-889.
- Montemagni, S., M. Wieling, B. de Jonge and J. Nerbonne (2013). Synchronic Patterns of Tuscan Phonetic Variation and Diachronic Change: Evidence from a Dialectometric Study. *Literary and Linguistic Computing*, 28(1), 157-172.
- Montemagni, S. (2007). Patterns of phonetic variation in Tuscany: using dialectometric techniques on multi-level representations of dialectal data. In P. Osenova et al. (eds.), *Proceedings of the Workshop on Computational Phonology at RANLP-2007*, pp. 49-60.
- Montemagni, S. (2008). Analisi linguistico-computazionali del corpus dialettale dell'Atlante Lessicale Toscano. Primi risultati sul rapporto toscano-italiano. In A. Nesi and N. Maraschio (eds.), *Discorsi di lingua e letteratura italiana per Teresa Poggi Salani* (Strumenti di filologia e critica, vol. 3), Pisa, Pacini, pp. 247-260.
- Nerbonne, J., W. Heeringa, E. van den Hout, P. van de Kooi, S. Otten, W. van de Vis (1996). Phonetic Distance between Dutch Dialects. In G. Durieux, W. Daelemans, S. Gillis (eds.), *Proceedings of the Sixth CLIN Meeting (Antwerp, Centre for Dutch Language and Speech, UIA)*, pp. 185-202.
- Nerbonne, J. and P. Kleiweg (2007). Toward a dialectological yardstick. *Journal of Quantitative Linguistics*, 14(2), 148-167.
- Nerbonne, J., (2003). Linguistic Variation and Computation. In Proceedings of the 10th Meeting of the European Chapter of the Association for Computational Linguistics, 15-17 April 2003, Budapest, Hungary, pp. 3-10
- Nerbonne, J. (2009). Data-driven dialectology. *Language and Linguistics Compass*, 3(1):, 175-198.
- Nerbonne, J. (2010). Measuring the diffusion of linguistic change. *Philosophical Transactions of the Royal Society B: Biological Sciences*, 365, 3821-3828.

- Pagel, M., Q. Atkinson and A. Meade (2007). Frequency of word-use predicts rates of lexical evolution throughout Indo-European history. *Nature*, 449, 717-720.
- Poggi Salani, T. (1978). Dialetto e lingua a confronto. In G. Giacomelli et al., Atlante lessicale toscano - Note al questionario, Firenze, Facoltà di Lettere e Filosofia, pp. 51-65.
- Seguy, J. (1973). La dialectométrie dans l'atlas linguistique de gascogne. *Revue de Linguistique Romane*, 37(145), 1–24.
- Tagliamonte, S.A. and R.H. Baayen (2012). Models, forests, and trees of York English: *Was/were* variation as a case study for statistical practice. *Language Variation and Change*, 24(2), 135-178.
- Valls, E., M. Wieling and J. Nerbonne (2013). Linguistic Advergence and Divergence in Northwestern Catalan: A Dialectometric Investigation of Dialect Leveling and Border Effects. *Literary and Linguistic Computing*, 28(1), 119-146.
- Wattenmaker, W.D. and E.J. Shoben (1987). Context and the recallability of concrete and abstract sentences. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 13(1), 140-150.
- Wieling, M., J. Nerbonne and R.H. Baayen (2011). Quantitative Social Dialectology: Explaining Linguistic Variation Socially and Geographically. *PLOS ONE*, 6(9), e23613.
- Wieling, M. and J. Nerbonne (2010). Hierarchical spectral partitioning of bipartite graphs to cluster dialects and identify distinguishing features. *Proceedings of the 2010 Workshop on Graph-based Methods for Natural Language Processing*, ACL, Uppsala, Sweden, 16 July 2010, pp. 33-41.
- Wieling, M. and J. Nerbonne (2011). Bipartite spectral graph partitioning for clustering dialect varieties and detecting their linguistic features. *Computer Speech and Language*, 25(3), pp. 700-715.
- Wieling, M., R.G. Shackleton and J. Nerbonne (2013). Analyzing phonetic variation in the traditional English dialects: Simultaneously clustering dialects and phonetic features. *Literary and Linguistic Computing*, 28(1), 31-41.
- Wieling, M. (2012). A Quantitative Approach to Social and Geographical Dialect Variation. PhD Thesis, University of Groningen.
- Wood, S. (2003). Thin plate regression splines. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 65(1), 95-114.

Wood, S. (2006). *Generalized additive models: an introduction with R*. Chapman & Hall/CRC.

Appendix: function call fitting the generalized additive model

```
library(mgcv)
```

```
# random intercepts and slopes are denoted by s(...,bs="re")
model = bam (UnequalToStandardItalian ~
CommunitySize.log + MaleGender + FarmerProfession +
EducationLevel.log +
te(Longitude,Latitude,ConceptFreqeuncy,SpeakerYearBirth,d=c(2,1,1)) +
s(Speaker,bs="re") + s(Location,bs="re") + s(Concept,bs="re") +
s(Word,YearOfRecording,bs="re") + s(Word,CommunitySize.log,bs="re") +
s(Word,AverageCommunityIncome.log,bs="re") +
s(Word,FarmerProfession,bs="re") +
s(Word,FarmerProfession,bs="re") +
s(Word,EaucationLevel.log,bs="re") + s(Word,MaleGender,bs="re"),
data=lexdst, family="binomial")
```

```
summary(model)
```